

# The differential transform method for solving random differential models

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## Main Problem

$$\begin{cases} \dot{X}(t) = P_n(t)X(t) + B(t); & P_n(t) := \sum_{i=0}^n a_i t^i, \quad t \in T, \\ X(0) = X_0 \end{cases} \quad (1)$$

where  $B(t)$  is a stochastic process and  $a_i, X_0$  are random variables.

- Construction of a solution of the form:

$$X(t) = \sum_{k=0}^{\infty} \dot{X}(k)(t-t_0)^k, \quad X: T \times \Omega \rightarrow \mathbb{R}, \quad T \subset \mathbb{R},$$

by means of Differential Transform Method.

- Numerical example and main statistical properties (expectation and standard deviation)

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- A real random variable  $X$  defined on  $(\Omega, \mathcal{F}, P)$  is called of order  $p$  ( $p$ -r.v.), if

$$E[|X|^p] < \infty, \quad p \geq 1, \quad (L_p, \|X\|_p := (E[|X|^p])^{1/p}); \quad p = 2$$

- Let  $\{X_n : n \geq 0\}$  be a sequence of  $p$ -r.v.'s. We say that it is convergent in the  $p$ -th mean to the  $p$ -r.v.  $X \in L_p$ , if

$$\lim_{n \rightarrow +\infty} \|X_n - X\|_p = 0.$$

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- If  $E[|X(t)|^p] < +\infty$  for all  $t \in T$ , then it is called a stochastic process of order  $p$  ( $p$ -s.p.)
- If there exists a stochastic process  $\frac{dX(t)}{dt}$  of order  $p$ , such that  $\left\| \frac{X(t+h) - X(t)}{h} - \frac{dX(t)}{dt} \right\|_p \rightarrow 0$  as  $h \rightarrow 0$ , then we say that  $\{X(t) : t \in T\}$  is  $p$ -th mean differentiable at  $t \in T$ .



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The random differential transform of the process  $X(t)$  is defined as

$$\hat{X}(k) = \frac{1}{k!} \left[ \frac{d^k (X(t))}{dt^k} \right]_{t=t_0}, \quad (2)$$

where  $\hat{X}$  is the transformed s.p and  $\frac{d}{dt}$  denotes de m.s. derivative.  
The inverse transform of  $\hat{X}$  is defined as

$$X(t) = \sum_{k=0}^{\infty} \hat{X}(k)(t - t_0)^k. \quad (3)$$

Table: Mathematical operations

Original s.p.	Transformed process <sup>1</sup>
$X(t) = Y(t) \pm Z(t)$	$\hat{X}(k) = \hat{Y}(k) \pm \hat{Z}(k)$
$X(t) = \lambda Y(t)$	$\hat{X}(k) = \lambda \hat{Y}(k)$
$Y(t) = \frac{dX(t)}{dt}$	$\hat{Y}(k) = (k+1)\hat{X}(k+1)$
$X(t) = Y(t)Z(t)$	$\hat{X}(k) = \sum_{r=0}^k \hat{Y}(r)\hat{Z}(k-r)$

<sup>1</sup>L. Villafuerte, C.A. Braumann, J.C. Cortés, L. Jódar, Random differential operational calculus: Theory and applications, Comput. Math. Appl. 59 (2010) 115–125.

$$\dot{X}(t) = P_n(t)X(t) + B(t)$$

- A p-s.p.  $\{X(t) : |t| < c\}$  is  **$p$ -th mean analytic on  $|t| < c$**  if it can be expanded in the  $p$ -th mean convergent Taylor series

$$X(t) = \sum_{n=0}^{\infty} (t - t_0)X^{(n)}(t_0)/n!$$

- Let  $\{H(t) : |t| < c\}$  be a  $p$ -th analytic s.p. given by

$$H(t) = \sum_{k=0}^{\infty} H_k t^k, \quad H_k = \frac{H^{(k)}(0)}{k!}, \quad |t| < c,$$

where the derivatives are considered in the  $p$ -th sense. Then there exists  $M > 0$  such that  $\|H_k\|_p \leq \frac{M}{\rho^k}$ ,  $0 < \rho < c$ ,  $\forall k \geq 0$

# Application to our problem- $X(t) = \sum_{k=0}^{\infty} \hat{X}(k)(t - t_0)^k$ .

- $\dot{X}(t) = P_n(t)X(t) + B(t)$ ;  $P_n(t) := \sum_{i=0}^n a_i t^i$

$$(k+1)\hat{X}(k+1) = \sum_{r=0}^k \hat{P}_n(r)\hat{X}(k-r) + \hat{B}(k) \quad (4)$$

$$\hat{P}_n(r) = \begin{cases} a_r, & 0 \leq r \leq n, \\ 0, & r > n. \end{cases} \quad (5)$$

Thus equation (??) becomes

$$(k+1)\hat{X}(k+1) = \sum_{r=0}^n a_r \hat{X}(k-r) + \hat{B}(k), \quad k = n, n+1, \dots \quad (6)$$

for which the r.v's  $\hat{X}(0), \dots, \hat{X}(n)$  have to be computed from (??).

$$(k+1)\hat{X}(k+1) = \sum_{r=0}^k \hat{P}_n(r)\hat{X}(k-r) + \hat{B}(k)$$

$$\vec{Z}(k) := (Z_1(k), Z_2(k), \dots, Z_n(k), Z_{n+1}(k))^T =$$

$$(\hat{X}(k-n), \hat{X}(k-n+1), \dots, \hat{X}(k-1), \hat{X}(k))^T,$$

$$\vec{Z}(k+1) = A(k)\vec{Z}(k) + \vec{H}(k) \quad (7)$$

where

$$A(k) = \begin{pmatrix} 0 & 1 & 0 & 0 & \cdots & 0 \\ 0 & 0 & 1 & 0 & \cdots & 0 \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & 0 & \cdots & 1 \\ \frac{a_n}{k+1} & \frac{a_{n-1}}{k+1} & \frac{a_{n-2}}{k+1} & \cdots & \cdots & \frac{a_0}{k+1} \end{pmatrix}, \quad \vec{H}(k) = \begin{pmatrix} 0 \\ 0 \\ \vdots \\ 0 \\ \frac{\hat{B}(k)}{k+1} \end{pmatrix}$$

# Conditions for the m.s. convergence of

$$X(t) = \sum_{k=0}^{\infty} \hat{X}(k)(t - t_0)^k$$

$$\vec{Z}(k) = \left( \prod_{i=n}^{k-1} A(i) \right) \vec{Z}(n) + \sum_{r=n}^{k-1} \left( \prod_{i=r+1}^{k-1} A(i) \right) \vec{H}(r) \quad (8)$$

for  $k = n, n+1, \dots$  Here

$$\prod_{i=n_0}^m A(i) = \begin{cases} A(m)A(m-1)\cdots A(n_0) & \text{if } m \geq n_0, \\ 1 & \text{otherwise.} \end{cases} \quad (9)$$

Because of  $\|\vec{Z}(k)\|_{4,v} \geq \|\hat{X}(k)\|_4$ ;  $(L_p^m, \|\vec{X}\|_{p,v} := \max \|X_j\|_p)$

$$\sum_{k=0}^{\infty} \|\vec{Z}(k)\|_{4,v} (t - t_0)^k < \infty \implies \sum_{k=0}^{\infty} \|\hat{X}(k)\|_4 (t - t_0)^k < \infty$$

$$\vec{Z}(k) = \left( \prod_{i=n}^{k-1} A(i) \right) \vec{Z}(n) + \sum_{r=n}^{k-1} \left( \prod_{i=r+1}^{k-1} A(i) \right) \vec{H}(r)$$

- Computing the p norm:

$$\left\| \vec{Z}(k) \right\|_{p,v} \leq \left\| \left( \prod_{i=n}^{k-1} A(i) \right) \vec{Z}(n) \right\|_{p,v} + \sum_{r=n}^{k-1} \left\| \left( \prod_{i=r+1}^{k-1} A(i) \right) \vec{H}(r) \right\|_{p,v}$$

- Using that  $\|A\|_{p,m} = \sum_{i,j} \|A_{ij}\|_p$

$$E[|XY|] \leq \|X\|_2 \|Y\|_2 \quad \implies \quad \|XY\|_p \leq \|X\|_{2p} \|Y\|_{2p}$$

and

$$\left\| \prod_{i=1}^m X_i \right\|_p \leq \prod_{i=1}^m \left( \|X_i\|_p \right)^{\frac{1}{m}}, \quad m \geq 1, \quad E[(X_i)^{mp}] < \infty, \quad \forall i.^2$$

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<sup>2</sup>J.-C. Cotés et al, Solving the random Legendre differential equation: Mean square power series solution and its statistical functions, Computers and Mathematics with Applications 61 (2011): 2782 - 2792



$$\left\| \vec{Z}(k) \right\|_{p,v} \leq \left\| \left( \prod_{i=n}^{k-1} A(i) \right) \vec{Z}(n) \right\|_{p,v} + \sum_{r=n}^{k-1} \left\| \left( \prod_{i=r+1}^{k-1} A(i) \right) \vec{H}(r) \right\|_{p,v}$$

- It follows

$$\left\| \prod_{i=n}^{k-1} A(i) \right\|_{2p,m} \leq \sum_{r,s,s_1,s_2,\dots,s_{k-n}=1}^{n+1} \left( \left\| (A_{rs_1}(k-1))^{k-n} \right\|_{2p} \right)^{\frac{1}{k-n}} \cdots \left( \left\| (A_{s_{k-n-1}s}(n))^{k-n} \right\|_{2p} \right)^{\frac{1}{k-n}}$$

- Now:

$$\left( \left\| (A_{rs}(l))^{k-n} \right\|_{2p} \right)^{\frac{1}{k-n}} = \begin{cases} 0 & \text{or,} \\ 1 & \text{or,} \\ \left( \mathbb{E} \left[ \left( \frac{a_j}{l+1} \right)^{2p(k-n)} \right] \right)^{\frac{1}{2p(k-n)}} & \end{cases} \quad (10)$$

- Recalling:

$$\|\vec{Z}(k)\|_{p,v} \leq \left\| \left( \prod_{i=n}^{k-1} A(i) \right) \vec{Z}(n) \right\|_{p,v} + \sum_{r=n}^{k-1} \left\| \left( \prod_{i=r+1}^{k-1} A(i) \right) \vec{H}(r) \right\|_{p,v}$$

- and  $\vec{H}(r) = (0, \dots, 0, \frac{\hat{B}(k)}{k+1})^T$

$$\|\vec{H}(r)\|_{2p,v} = \left\| \frac{\hat{B}(r)}{r+1} \right\|_{2p} \leq \frac{M_2}{(r+1)\rho^r}, \quad (11)$$

for  $0 < \rho < c$ .

## Theorem

Consider the problem

$$\begin{cases} \dot{X}(t) = P_n(t)X(t) + B(t); & P_n(t) := \sum_{i=0}^n a_i t^i, \quad t \in T, \\ X(0) = X_0 \end{cases} \quad (12)$$

Assume that the random variables  $a_j$  satisfy the condition

$$E[|a_j^m|] \leq KM^m < \infty, \quad \forall m \geq 0, \quad j = 0, 1, \dots, n$$

$X_0$  is 4-r.v and  $B(t)$  is an 2-th analytic stochastic process. Then there exists a solution of the form

$$X(t) = \sum_{k=0}^{\infty} \hat{X}(k)t^k; \quad \hat{X}(k+1) = \frac{1}{k+1} \left( \sum_{r=0}^k \hat{P}_n(r)\hat{X}(k-r) + \hat{B}(k) \right)$$

$$|t| < c = \min(\rho, \frac{1}{n+1})$$

# Approximations of the mean and standard deviation functions

- 1 The truncated process and its main moments

$$X_N(t) = \sum_{k=0}^N \hat{X}(k)t^k; \quad t_0 = 0.$$

$$E[X_N(t)] = \sum_{k=0}^N E[\hat{X}(k)] t^k$$

$$E[(X_N(t))^2] = \sum_{k=0}^N E[(\hat{X}(k))^2] t^{2k} + 2 \sum_{k=1}^N \sum_{l=0}^{k-1} E[(\hat{X}(k)\hat{X}(l))] t^{k+l}$$

- 2 Advantages of the mean square convergence

$$X_N(t) \xrightarrow{m.s.} X(t) \Rightarrow E[X_N(t)] \rightarrow E[X(t)]$$

$$X_N(t) \xrightarrow{m.s.} X(t) \Rightarrow \text{Var}[X_N(t)] \rightarrow \text{Var}[X(t)]$$

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## Example

Consider:

$$\begin{cases} \dot{X}(t) = atX(t) + e^{-bt}; & t \in [0, 1/2], \\ X(0) = X_0 \end{cases} \quad (13)$$

where  $a$  is a standard normal r.v. ( $a \sim N(0,1)$ ),  $b$  is a exponential r.v. with parameter  $\lambda = 1$  ( $b \sim \text{Exp}[\lambda = 1]$ ), and  $X_0$  is a Beta r.v. with parameters  $\alpha = 2$  and  $\beta = 3$ , ( $X_0 \sim \text{Be}(\alpha = 2, \beta = 3)$ )

# Numerical results

$t$	$E[X(t)]$	$E[X_N(t)]$	$\sigma[X(t)]$	$\sigma[X_N(t)]$
0.00	0.250000	0.25000 0	0.193649	0.193389
0.10	0.345314	0.345314	0.193711	0.193448
0.20	0.432392	0.432392	0.194511	0.194243
0.30	0.512766	0.512766	0.197620	0.197344
0.40	0.587884	0.587884	0.205386	0.205102
0.50	0.659248	0.659248	0.220871	0.220580

Table: Comparison of the expectation and standard function

# THANK YOU!