

UNIFORM CONVERGENCE OF KERNEL ESTIMATORS  
WITH RANDOM BANDWIDTH

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Abstract

Standard kernel estimators do not converge to the true distribution uniformly. A consequence is that no inequality like Dvoretzky-Kiefer-Wolfowitz one can be constructed, and as a result it is impossible to answer the question how many observations are needed to guarantee a prescribed level of accuracy of the estimator. A remedy is to adapt the bandwidth to the sample at hand.

*Key Words:* kernel estimators, asymptotics, Glivenko-Cantelli theorem, Dvoretzky-Kiefer-Wolfowitz inequality, bandwidth, adaptive estimators, uniform limit laws

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1. GLIVENKO-CANTELLI THEOREM AND DVORETZKY-KIEFER-WOLFOWITZ INEQUALITY. Let  $X_1, X_2, \dots, X_n$  be a sample from an (unknown) distribution  $F \in \mathcal{F}$  where  $\mathcal{F}$  is the class of all continuous distribution functions.

The version of the Glivenko-Cantelli theorem in the form to be exploited below states that

$$(GCT) \quad (\forall \varepsilon)(\forall \eta)(\exists N)(\forall n \geq N)(\forall F \in \mathcal{F}) \quad P\left\{ \sup_{x \in \mathbf{R}^1} |F_n(x) - F(x)| \geq \varepsilon \right\} \leq \eta$$

where

$$F_n(x) = \frac{1}{n} \sum_{j=1}^n 1_{(-\infty, x]}(X_j).$$

The theorem is effective in the sense that for every  $\varepsilon > 0$  and for every  $\eta > 0$  one can effectively calculate  $N = N(\varepsilon, \eta)$ . That can be done by the following version of Dvoretzky-Kiefer-Wolfowitz inequality (Massart 1990)

$$(*) \quad P\left\{ \sup_{x \in \mathbf{R}^1} |F_n(x) - F(x)| \geq \varepsilon \right\} \leq 2e^{-2n\varepsilon^2}.$$

Due to the above, *GCT* together with  $(*)$  give us a genuinely statistical tool; if all that a statistician knows is that an unknown distribution  $F$  belongs to  $\mathcal{F}$ , he is able to make a precise inference about  $F$  (testing hypotheses or constructing confidence intervals).

2. KERNEL ESTIMATORS. The standard kernel density estimator is of the form (e.g. Wegman 2006)

$$\hat{f}_n(x) = \frac{1}{n} \sum_{j=1}^n \frac{1}{h_n} k\left(\frac{x - X_j}{h_n}\right)$$

with appropriate  $h_n, n = 1, 2, \dots$ . We shall consider kernel distribution estimator in its classical form

$$\hat{F}_n(x) = \frac{1}{n} \sum_{j=1}^n K\left(\frac{x - X_j}{h_n}\right)$$

where  $K(x) = \int_{-\infty}^x k(t)dt$ , and we shall show that (GCT) does not hold if  $F_n$  is replaced by  $\hat{F}_n$ , i.e. that the following is true

$$(\exists \varepsilon)(\exists \eta)(\forall N)(\exists n \geq N)(\exists F \in \mathcal{F}) \quad P\left\{\sup_{x \in \mathbf{R}^1} |\hat{F}_n(x) - F(x)| \geq \varepsilon\right\} \geq \eta.$$

Obviously it is enough to demonstrate that

$$(\dagger) \quad (\exists \varepsilon)(\exists \eta)(\forall n)(\exists F \in \mathcal{F}) \quad P\{\hat{F}_n(0) > F(0) + \varepsilon\} \geq \eta.$$

Concerning the kernel  $K$ , only the following assumptions are relevant: 1)  $0 < K(0) < 1$  and 2)  $K^{-1}(t) < 0$  for some  $t \in (0, K(0))$ . Concerning the sequence  $(h_n, n = 1, 2, \dots)$  we assume that  $h_n > 0, n = 1, 2, \dots$

Take  $\varepsilon \in (0, t)$  and  $\eta \in (t - \varepsilon, 1)$ . Given  $\varepsilon, \eta$ , and  $n$ , take  $F$  such that  $F(0) = t - \varepsilon$  and  $F(-h_n K^{-1}(t)) = P\{X_j < -h_n K^{-1}(t)\} > \eta^{1/n}$ . Then

$$P\left\{K\left(-\frac{X_j}{h_n}\right) > t\right\} > \eta^{1/n}$$

and due to the fact that

$$\bigcap_{j=1}^n \left\{K\left(-\frac{X_j}{h_n}\right) > F(0) + \varepsilon\right\} \subset \left\{\frac{1}{n} \sum_{j=1}^n K\left(-\frac{X_j}{h_n}\right) > F(0) + \varepsilon\right\}$$

we have

$$P\left\{\frac{1}{n} \sum_{j=1}^n K\left(-\frac{X_j}{h_n}\right) > t\right\} = P\left\{\frac{1}{n} \sum_{j=1}^n K\left(-\frac{X_j}{h_n}\right) > F(0) + \varepsilon\right\} > \eta;$$

hence  $(\dagger)$ .

It follows that for classical kernel estimators no inequality like (\*) can be obtained which makes the estimators of a doubtful usefulness for statistical applications.

3. RANDOM BANDWIDTH. A remedy is as follows. Let  $X_{1:n} \leq X_{2:n} \leq \dots \leq X_{n:n}$  be order statistics from the sample  $X_1, X_2, \dots, X_n$ . Define

$$H_n = \min\{X_{j:n} - X_{j-1:n}, j = 2, 3, \dots, n\}.$$

Define the kernel estimator

$$\tilde{F}_n(x) = \frac{1}{n} \sum_{j=1}^n K\left(\frac{x - X_j}{H_n}\right)$$

where for  $K$  we assume:

$$K(t) = \begin{cases} 0, & \text{for } t \leq -\frac{1}{2}, \\ 1, & \text{for } t \geq \frac{1}{2}, \end{cases}$$

$K(0) = \frac{1}{2}$ ,  $K(t)$  continuous and nondecreasing in  $(-\frac{1}{2}, \frac{1}{2})$ .

Now, for  $k = 1, 2, \dots, n$  we have  $|\tilde{F}_n(X_{k:n}) - F_n(X_{k:n})| \leq \frac{1}{2n}$ . Kernel estimator  $\tilde{F}_n(x)$  is continuous and increasing, empirical distribution function  $F_n(x)$  is a step function, and in consequence  $|\tilde{F}_n(x) - F_n(x)| \leq \frac{1}{2n}$  for all  $x \in (-\infty, \infty)$ . By the triangle inequality

$$|\tilde{F}_n(x) - F(x)| \leq |F_n(x) - F(x)| + \frac{1}{2n}$$

we obtain

$$P\left\{\sup_{x \in \mathbf{R}^1} |\tilde{F}_n(x) - F(x)| \geq \varepsilon\right\} \leq P\left\{\sup_{x \in \mathbf{R}^1} |F_n(x) - F(x)| + \frac{1}{2n} \geq \varepsilon\right\}$$

and hence, by (\*) we have

$$(**) \quad P\left\{\sup_{x \in \mathbf{R}^1} |\tilde{F}_n(x) - F(x)| \geq \varepsilon\right\} \leq 2e^{-2n(\varepsilon - 1/2n)^2}, \quad n > \frac{1}{2\varepsilon}$$

which enables us to calculate  $N = N(\varepsilon, \eta)$  that guarantees the prescribed accuracy of the kernel estimator  $\tilde{F}_n(x)$ .

A COMMENT. Observe that the smallest  $N = N(\varepsilon, \eta)$  that guarantees the prescribed accuracy is somewhat greater for kernel estimator  $\tilde{F}_n$  than that for crude empirical step function  $F_n$ . For example,  $N(0.1, 0.1) = 150$  for  $F_n$  and  $= 160$  for  $\tilde{F}_n$ ;  $N(0.01, 0.01) = 26,492$  for  $F_n$  and  $= 26,592$  for  $\tilde{F}_n$ . Another disadvantage of kernel smoothing has been discovered by Hjort and Walker (2001): "kernel density estimator with optimal bandwidth lies outside any confidence interval, around the empirical distribution function, with probability tending to 1 as the sample size increases". Perhaps a reason is that smoothing adds to observations something which is rather arbitrarily chosen and which may spoil the inference.

A GENERALIZATION. Inequality (\*\*) holds for every smoothed nondecreasing distribution function  $\tilde{F}_n(x)$  that satisfies  $|\tilde{F}_n(X_{k:n}) - F_n(X_{k:n})| \leq \frac{1}{2n}$ ,  $k = 1, 2, \dots, n$ .

## References

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