

**AN ABC OF PORTFOLIO CHOICE:  
ASSET ALLOCATION  
WITH BANKRUPTCY AND CONTAGION**

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## The Problem

- Investors like to invest in all asset classes including corporate bonds.
- Especially in times of low interest rates corporate bonds are sometimes considered as a good investment alternative.
- However, there has been not much work on portfolio problems with default risk.
- We try to fill this gap by considering such a problem with contagion and bankruptcy.

## Example for Contagion Effects

- On October 8, 2005 the auto parts maker Delphi Corp. filed for Chapter 11 bankruptcy protection.
- On December 12, 2005 the rating agency S&P cut GM's corporate credit rating to B, five steps below investment grade.
- On March 3, 2006 the auto parts maker Dana Corp. filed for bankruptcy protection defaulting on \$2.5 billion of debt.

## Related Work

### Portfolio Optimization in Firm Value Models

- Korn, K. (2003)
- K., Steffensen (2004)

### Portfolio Optimization in a Standard Reduced-from Model

- K., Steffensen (2005)

### Portfolio Optimization using the JLY-Assumption

- Jarrow, Lando, Yu (2005): Conditionally Diversifiable Default Risk
- Hou (2003)

### Pricing Defaultable Claims using Markov Chains

- K., Steffensen (2006)
- Schonbucher (2006)

## JLY-Assumption: Conditionally Diversifiable Default Risk

Roughly speaking, **JLY prove the following**: If there are

- **infinitely** many firms and
- individual default-timing risk can be **diversified**

then default-timing risk premiums are zero.

Imposing this assumption in a portfolio problem means that one **gets rid off all jump processes**.

→ Such a problem is similar to a problem with stochastic interest rates.

(See Sørensen (1998), Korn, K. (2001), Munk, Sørensen (2004))

## Is the JLY-assumption Supported by Empirical Evidence?

- After accounting for tax and liquidity, **Driessen (2003)** reports an average default-timing risk premium across his data of 1.89.
- This result is roughly in line with the results of **Berndt, Douglas, Duffie, Ferguson, Schranz (2004)**.
- **Yu (2002)** demonstrates how a positive default-timing risk premium can be used to resolve some empirical puzzles in the credit risk literature.

**These studies do not support the JLY-assumption.**

⇒ Default-timing risk should not be ignored in portfolio problems.

## Agenda

1. **Portfolio Problem**
  - Merton's Results
  - Markov Chain Framework
2. **Solution**
  - First-order Conditions
  - Representation of the Investor's Indirect Utility
3. **Applications**
  - Reorganization vs. Liquidation
  - Contagion

## Merton's Results

An Investor allocates his wealth between a stock and a risky bond:

$$\begin{aligned}dS(t) &= S(t)[\alpha dt + \sigma dW(t)], & S(0) &= s_0, \\dB(t) &= B(t)[r dt - dN(t)], & B(0) &= b_0.\end{aligned}$$

$r$  is the short rate and  $N$  is a Poisson process with intensity  $\lambda > 0$ .

**Implication:** Recovery rate is zero for the risky bond.



## Merton's Results

Merton's HJB for the indirect utility  $G$  ( $\pi$ : stock proportion,  $x$ : wealth):

$$0 = \sup_{\pi} \{G_t(t, x) + (r + (\alpha - r)\pi)G_x(t, x) + 0.5x^2\pi^2x^2G_{xx}(t, x) + \lambda[G(t, \pi x) - G(t, x)]\}, \quad G(T, x) = \frac{1}{\gamma}x^\gamma.$$

**Problem:** In case of a default, the **bond is wiped out**, but Merton implicitly assumes that the investor can still invest in bonds.

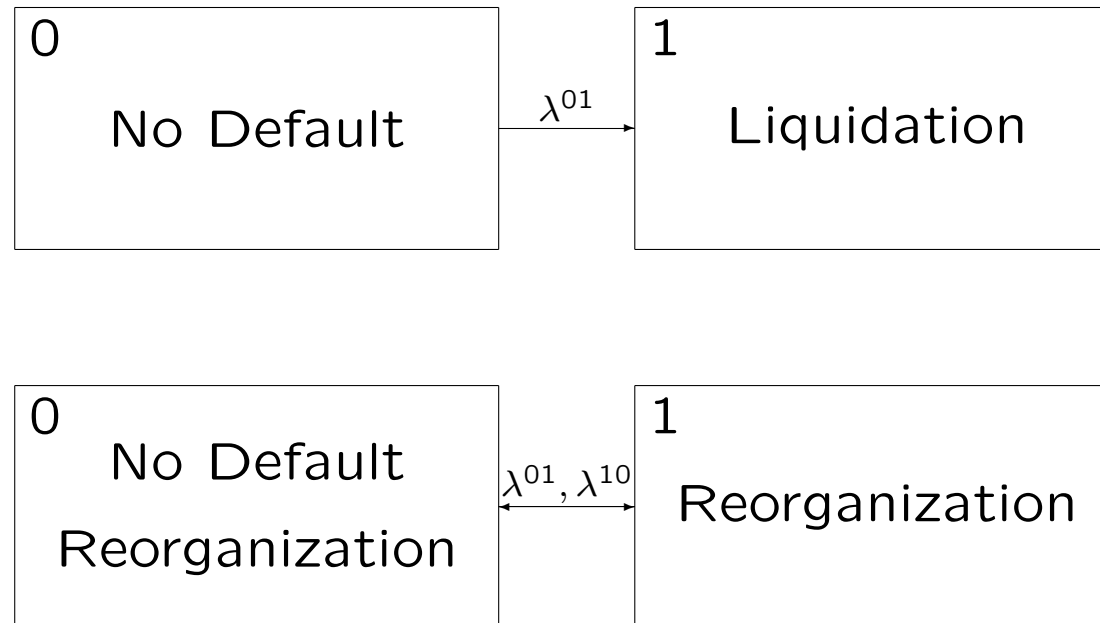
$$0 = \sup_{\pi} \{G_t(t, x) + (r + (\alpha - r)\pi)G_x(t, x) + 0.5x^2\pi^2x^2G_{xx}(t, x) + \lambda[\tilde{G}(t, \pi x) - G(t, x)]\}.$$

$$\tilde{G}(t, x) = \mathbb{E}^{t,x} \left[ \frac{1}{\gamma} \left( \frac{x_t}{s_t} \right)^\gamma S(T)^\gamma \right] = \frac{1}{\gamma}x^\gamma \exp(\gamma(\alpha - 0.5(1 - \gamma)\sigma^2)(T - t)),$$

In general,  $G \neq \tilde{G}$

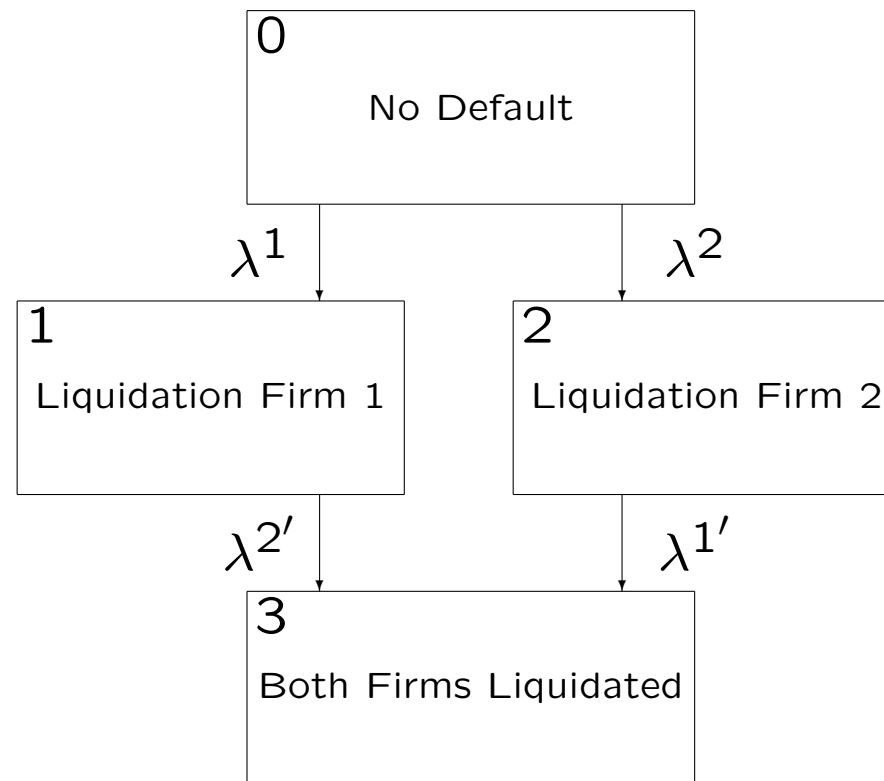
## Why are Markov Chains Useful?

### Models for a Single Firm: Liquidation vs. Reorganization



**Classical Models like Jarrow/Turnbull (1995) or Duffie/Singleton (1999) can be embedded in a Markov Chain Framework.**

## Why are Markov Chains Useful?



**The Contagion Model by Jarrow/Yu (2001) can be embedded in a Markov Chain Framework.**

## Standing Assumption

To rule out arbitrage, we assume that an **equivalent martingale measure  $Q$  exists** under which discounted asset prices are (local) martingales.

See Harrison and Kreps (1979) and Delbaen and Schachermayer (1999) for the essential equivalence of the existence of such a measure and the absence of arbitrage.

## Markov Chain Framework

$\mathcal{J} = \{0, \dots, J\}$ : **finite set of states** characterizing the market for corporate claims

$Z(t)$ : **state at time**  $t \in [0, T^*]$  (RCLL process)

$N = (N^k)_{k \in \mathcal{J}}$ :  $J$ -dimensional **counting process** counting the number of transitions into the states, i.e.

$$N^k(t) = \#\{s | s \in (0, t], Z(s-) \neq k, Z(s) = k\}.$$

$Y = (Y_1, \dots, Y_L)$ : **macroeconomic variables** with dynamics

$$dY_l(t) = \alpha_{Q,l}(t)dt + \beta_l(t)dW_Q(t),$$

where  $W_Q(t) = W(t) + \int_0^t \eta_y(s) ds$  is an  $L$ -dimensional standard Brownian increment under  $Q$ .

## Markov Chain Framework

**1st Assumption:** The processes  $N^K$  have  **$Q$ -intensities**  $\lambda_Q^{jk,n}(t, Y(t))$  such that

$$M_Q^k(t) = N^k(t) - \int_0^t \lambda_Q^{Z(s)k, N(s)}(s) ds$$

is a  $Q$ -martingale.

**2nd Assumption:** There exist strictly positive processes  $\eta^{jk,n}(t, Y(t))$  such that  $\lambda_Q^{jk,n} = \eta^{jk,n} \lambda^{jk,n}$ , where  $\lambda^{jk,n}$  denote the **physical intensities**, and

$$M^k(t) = N^k(t) - \int_0^t \lambda^{Z(s)k, N(s)}(s) ds$$

is a  $P$ -martingale.

## Markov Chain Framework

Consider a **corporate contingent claim** characterized by a payment stream

$$dA(t) = c^{Z(t)} dt + \sum_{k:k \neq Z(t-)} a^{Z(t-),k,N(t-)}(t) dN^k(t) + a^{Z(t),N(t)} d\varepsilon_T(t),$$

- $c^j$ : contingent coupon payment (continuously paid)
- $a^{jk,n}$ : payment upon transition from state  $j$  into state  $k$
- $a^{j,n}$ : final payment given state  $j$  at time  $T$
- $\varepsilon_T$ : Dirac mass at  $T$

### Example: Zero-coupon bond

- state 0: firm solvent
- state 1: firm in default
- $a^{01}$ : recovery payment upon default
- $a^0 = 1, a^1 = 0$ : state dependent final payments

## Markov Chain Framework

Under a risk-neutral measure, the **time- $t$  value** reads

$$\begin{aligned} B(t) &= \mathbb{E}_Q \left[ \int_{(t,T]} e^{-\int_t^s r(u) du} dA(s) \middle| \mathcal{F}_t \right] \\ &= \mathbb{E}_Q \left[ \int_{(t,T]} e^{-\int_t^s r(u) du} dA(s) \middle| Z(t), Y(t), N(t) \right], \end{aligned}$$

**Theorem (System of Pricing Equations)** There exist functions  $\mathcal{B}^{j,n}$ ,  $j \in \mathcal{J}$  and  $n \in \mathbb{N}^{J+1}$ , such that  $B(t) = \mathcal{B}^{Z(t),N(t)}(t, Y(t))$  and the functions  $\mathcal{B}^{j,n}$  satisfy the system of PDEs

$$\mathcal{B}_t^{j,n} = r\mathcal{B}^{j,n} - \alpha_Q \mathcal{B}_y^{j,n} - 0.5 \operatorname{tr}[\beta^T \mathcal{B}_{yy}^{j,n} \beta] - \sum_{k:k \neq j} \lambda_Q^{jk,n} \left( a^{jk,n} + \mathcal{B}^{k,n+1^k} - \mathcal{B}^{j,n} \right)$$

with terminal conditions  $\mathcal{B}^{j,n}(T, y) = a^{j,n}(y)$ ,  $j \in \mathcal{J}$  and  $n \in \mathbb{N}^{J+1}$ .



## Markov Chain Framework

**Proposition (Physical Return Dynamics)** Under the physical measure, the dynamics of the gains process  $\Gamma = B + A$  are given by

$$\begin{aligned} \frac{d\Gamma(t)}{B(t-)} &= (r(t) + \hat{\chi}^{Z(t),N(t)}(t))dt - D^{Z(t),N(t)}(t)\beta(t)dW(t) \\ &+ \sum_{k:k \neq Z(t-)} \tilde{L}^{Z(t-),k,N(t-)} dM^k(t) \end{aligned}$$

with

$$\hat{\chi}^{j,n} = - \sum_{k:k \neq j} \tilde{L}^{jk,n} (\eta^{jk,n} - 1) \lambda^{jk,n} - \beta D^{j,n} \eta_y,$$

$$D^{j,n} = -\mathcal{B}_y^{j,n} / \mathcal{B}^{j,n}:$$

state dependent duration,

$$\tilde{L}^{jk,n} = R^{jk,n} + L^{jk,n}:$$

actual loss

$$R^{jk,n} = a^{jk,n} / \mathcal{B}^{j,n}:$$

relative payment upon transition

$$L^{jk,n} = (\mathcal{B}^{k,n+1^k} - \mathcal{B}^{j,n}) / \mathcal{B}^{j,n}:$$

relative price jump.

## Portfolio Problem

**For simplicity:**  $Y = r$ , i.e. short rate is the only macro variable

### Money Market Account

$$dM(t) = M(t)r(t)dt, \quad M(0) = 1,$$

### Default-free zero-coupon bond

$$dP_f(t, T) = P_f(t, T)[(r(t) + \sigma_f(t)\eta_y(t))dt + \sigma_f(t)dW(t)], \quad P(T, T) = 1.$$

### Corporate zero-coupon bonds

$$\begin{aligned} \frac{d\Gamma(t)}{B(t-)} &= (r(t) + \chi^{Z(t), N(t)}(t))dt - D^{Z(t), N(t)}(t)\beta(t)dW(t) \\ &+ \sum_{k: k \neq Z(t-)} \tilde{L}^{Z(t-), k, N(t-)} dN^k(t) \end{aligned}$$

**Investor's time- $t$  wealth** is given by

$$X(t) = \varphi_f(t)P_f(t, T_f) + \sum_{i=1}^I \varphi_i(t)B_i(t, T_i) + \varphi_M(t)M(t),$$

For a self-financing strategy  $(\varphi_f, \varphi, \varphi_M)$ ,

$$dX(t) = \varphi_f(t)dP_f(t) + \sum_{i=1}^I \varphi_i(t-)d\Gamma_i(t, T_i) + \varphi_M(t)dM(t),$$

Setting  $\pi_0 = \varphi_f P_f / X$  and  $\pi_i = \varphi_i B_i / X$ , the **wealth equation** reads:

$$dX = X^- \left[ (r + \sum_{i=0}^I \pi_i \chi_i^{Z, N}) dt - \beta D dW + \sum_{i=1}^I \sum_{k \neq j} \pi_i^- \tilde{L}_i^{Z^-, N^-} dN^k \right],$$

where  $X^- := X(t-)$ ,  $Z^- = Z(t-)$ ,  $\pi_i^- := \pi_i(t-)$ ,  $\chi_0 = \chi_0^Z := \eta_r \sigma_f$ ,  
 $D := \sum_{i=0}^I \pi_i D_i^{Z, N}$  (portfolio duration).

0-th bond is identified with the default-free zero-coupon bond

## Portfolio Problem

The investor **maximizes expected utility** from intermediate consumption and terminal wealth at final time  $T$  w.r.t.

$$U(t, x) = \frac{1}{\gamma} \psi(t) x^{1-\gamma}, \quad \gamma < 1,$$

where  $\psi(t) = \psi(t, r(t))$  models the investor's time-preferences.

**Investor's indirect utility** (syn. value function) is defined by

$$G^{j,n}(t, x, r) = \sup_{\pi, D, c} \mathbb{E}_{j,n}^{t,x,r} \left[ \int_t^T \frac{1}{\gamma} \psi(s) x^{1-\gamma} c(s)^\gamma ds + \frac{1}{\gamma} \psi(T) x^{1-\gamma} (X(T))^\gamma \right]$$

## Hamilton-Jacobi-Bellman Equation

$$0 = \sup_{\pi, D, c} \left\{ A^{\pi, D, c} G^{j, n} + \frac{1}{\gamma} \psi^{1-\gamma} c^\gamma \right\}, \quad j \in \{0, \dots, J\},$$

with terminal condition  $G^{j, n}(T, x, r) = \frac{1}{\gamma} \psi(T)^{1-\gamma} x^\gamma$  and

$$\begin{aligned} A^{\pi, D, c} G^{j, n} = & G_t^{j, n} + \left\{ x(r - \sum_k \eta^{jk, n} \lambda^{jk, n} \sum_{i=1}^I \pi_i \tilde{L}_i^{jk, n} - \beta \eta_r D) - c \right\} G_x^{j, n} \\ & + \alpha G_r^{j, n} + 0.5 x^2 \beta^2 D^2 G_{xx}^{j, n} - x \beta^2 D G_{xr}^{j, n} + 0.5 \beta^2 G_{rr}^{j, n} \\ & + \sum_{k \neq j} \lambda^{jk, n} \left\{ G^{k, n+1^k}(t, x(1 + \sum_{i=1}^I \pi_i \tilde{L}_i^{jk, n}), r) - G^{j, n}(t, x, r) \right\}. \end{aligned}$$

We conjecture  $G^{j, n}(t, x, r) = \frac{1}{\gamma} x^\gamma f^{j, n}(t, r)^{1-\gamma}$  with  $f^{j, n}(T, r) = \psi(T)$

...

**Theorem (FOC) First-order conditions** for  $D$ ,  $\pi$ , and  $c$ :

$$\begin{aligned}
 D^{j,n*} &= -\frac{1}{1-\gamma} \frac{\eta_r}{\beta} - \frac{f_r^{j,n}}{f^{j,n}}, & \frac{c^{j,n*}}{x} &= \frac{\psi}{f^{j,n}} \\
 0 &= -\sum_{k \neq j} \tilde{L}_\nu^{jk,n} \eta^{jk,n} \lambda^{jk,n} \\
 &\quad + \sum_{k \neq j} \lambda^{jk,n} (1 + \sum_i \pi_i^{j,n*} \tilde{L}_i^{jk,n})^{\gamma-1} \left( \frac{f^{k,n+1^k}}{f^{j,n}} \right)^{1-\gamma} \tilde{L}_\nu^{jk,n},
 \end{aligned}$$

where the functions  $f^{j,n}$  satisfy the system of PDEs

$$\begin{aligned}
 0 &= \psi + f_t^{j,n} + \frac{\gamma}{1-\gamma} [r + 0.5 \frac{1}{1-\gamma} \eta_r^2 - \sum_{k \neq j} \eta^{jk,n} \lambda^{jk,n} \sum_{i=1}^I \pi_i^{j,n*} \tilde{L}_i^{jk,n}] f^{j,n} \\
 &\quad + (\alpha + \frac{\gamma}{1-\gamma} \beta \eta_r) f_r^{j,n} + 0.5 \beta^2 f_{rr}^{j,n} \\
 &\quad + \frac{1}{1-\gamma} f^{j,n} \sum_{k \neq j} \lambda^{jk,n} \left[ (1 + \sum_{i=1}^I \pi_i^{j,n*} \tilde{L}_i^{jk,n})^\gamma \left( \frac{f^{k,n+1^k}}{f^{j,n}} \right)^{1-\gamma} - 1 \right]
 \end{aligned}$$

with  $f^{j,n}(t, r) = 1$ ,  $j \in J$  and  $n \in IN_0^{J+1}$ . The **investor's indirect utility** is given by  $G^{j,n}(t, x, r) = \frac{1}{\gamma} x^\gamma f^{j,n}(t, r)^{1-\gamma}$ .

## Applications: Reorganization vs. Liquidation

Simple Problem: Money market account and one corporate bond

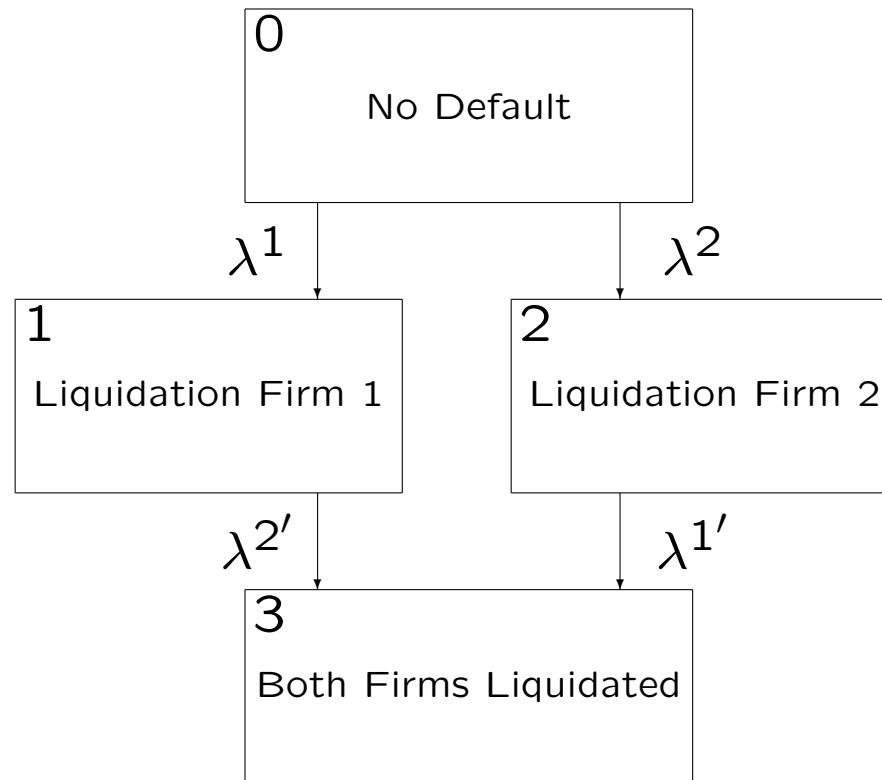
FOC: **Optimal bond demands** for economies with  
**Reorganization (Chapter 11)** and **Liquidation (Chapter 7)**:

$$\pi^* = \frac{1 - \eta^{\frac{1}{\gamma-1}}}{-L}, \quad \pi^* = \frac{1 - \eta^{\frac{1}{\gamma-1}} \frac{f^1}{f^0}}{-L}.$$

**Proposition** For positive default risk premiums, i.e.  $\eta - 1 \geq 0$ , we get  $f^0 \leq f^1$  if  $\gamma < 0$  and  $f^0 \geq f^1$  if  $\gamma > 0$

$\implies$  The **bond demand** is **smaller** in an economy where **the firm is liquidated**.

## Applications: Contagion



Contagion occurs if  $\lambda^{1'} > \lambda^1$  or  $\lambda^{2'} > \lambda^2$ .



## Applications: Contagion

**For simplicity:** Two corporate zeros with identical parameters, i.e.  $\lambda := \lambda^1 = \lambda^2$ ,  $\lambda' := \lambda^{1'} = \lambda^{2'}$ ,  $\eta := \eta^1 = \eta^2$ ,  $\eta' := \eta^{1'} = \eta^{2'}$ ,  $L_1^{01} = L_2^{02}$ ,  $L_1^{02} = L_2^{01}$ .

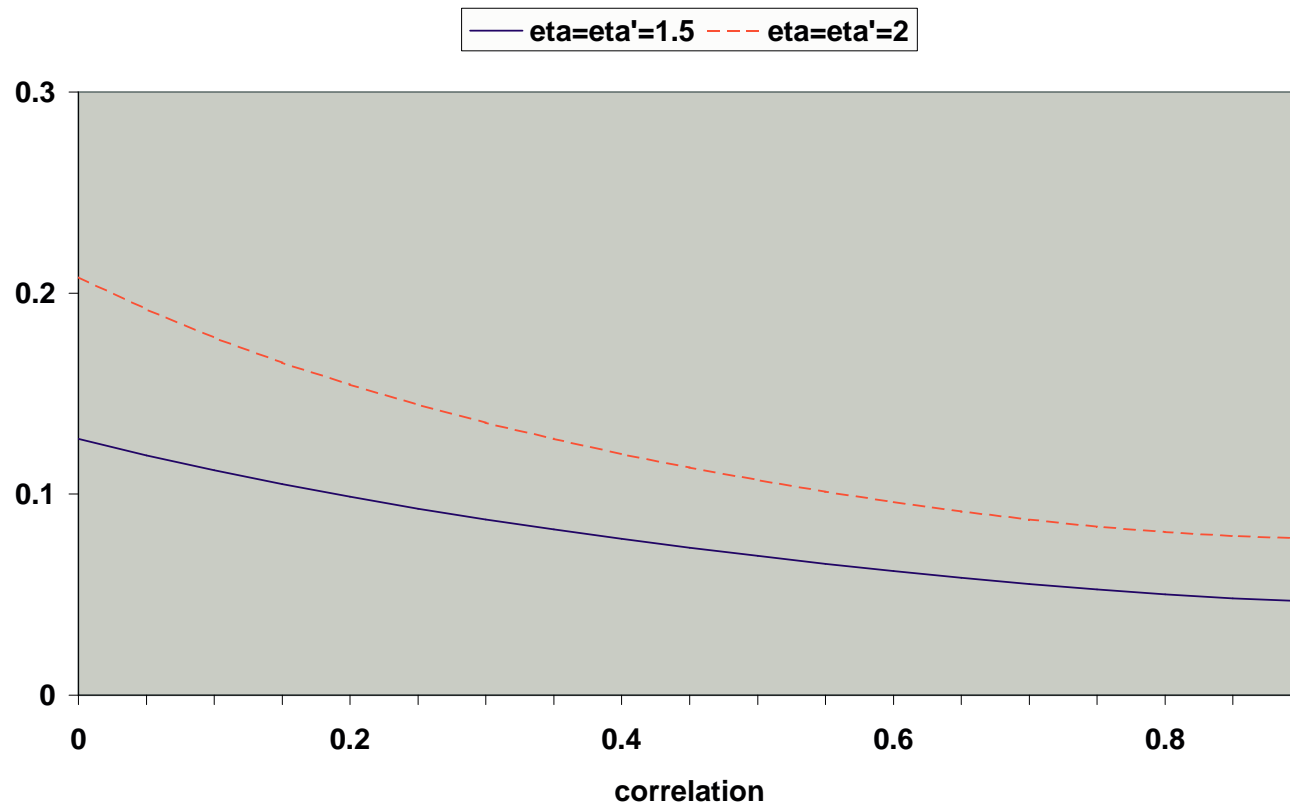
### Optimal demands

$$\pi_1^* = \frac{1 - \eta^{\frac{1}{\gamma-1}} \frac{f^1}{f^0}}{-(\tilde{L}_1^{01} + \tilde{L}_1^{02})}.$$

The effect of  $f^1/f^0$  on the optimal demands is small compared to the effect of  $\tilde{L}_1^{02}$ .

**Interpretation:** If we have a **positive contagion effect** (under the risk-neutral measure), i.e.  $\eta\lambda = \lambda_Q < \lambda'_Q = \eta'\lambda'$ , then  $\tilde{L}_1^{02} < 0$  and the **demand decreases**, whereas the opposite is true for a negative contagion effect (competition effect).

## Applications: Contagion



**Optimal Demands for Different Default Correlations**

## Conclusion

- We present a sophisticated framework covering various realistic features of defaultable bonds (e.g. Chapter 7 vs. Chapter 11)
- We are able to solve various involved portfolio problems with corporate bonds.
- In particular, we are able to *quantify* the impact of contagion on bond demands.
- This is important because contagion reduces the investor's ability to diversify his portfolio.
- Our approach can be applied to several other situations: ratings, default vs. bankruptcy, exogenous (re)issuing of bonds

$$\begin{aligned}
\rho_{1,2}(t, T) &= \frac{\text{Cov}_t(\mathbf{1}_{\{\tau_1 \leq T\}}, \mathbf{1}_{\{\tau_2 \leq T\}})}{\sqrt{\text{Var}_t(\mathbf{1}_{\{\tau_1 \leq T\}})} \sqrt{\text{Var}_t(\mathbf{1}_{\{\tau_2 \leq T\}})}} \\
&= \frac{p^{03} - (p^{01} + p^{03})(p^{02} + p^{03})}{\sqrt{(p^{01} + p^{03})(1 - p^{01} - p^{03})} \sqrt{(p^{02} + p^{03})(1 - p^{02} - p^{03})}}
\end{aligned}$$