



General Edgeworth  
expansions

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One-split branching random  
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Random trees

Mode and width

# General Edgeworth expansions with applications to profiles of random trees

Alexander Marynych

Taras Shevchenko National University of Kyiv  
and  
University of Münster

joint work with Z. Kabluchko (Münster, Germany) and  
H. Sulzbach (Birmingham, UK)



# Assumptions on profiles

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A random profile is an arbitrary sequence

$$\mathbb{L}_n = (\mathbb{L}_n(k))_{k \in \mathbb{Z}}, \quad n \in \mathbb{N}$$

of real-valued stochastic processes on  $\mathbb{Z}$ .



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of real-valued stochastic processes on  $\mathbb{Z}$ .

**Assumption A1:** There exists an open interval  $(\beta_-, \beta_+) \subset \mathbb{R}$  containing zero such that for every  $n \in \mathbb{N}$  and every  $\beta \in (\beta_-, \beta_+)$ ,

$$\sum_{k \in \mathbb{Z}} |\mathbb{L}_n(k)| e^{\beta k} < \infty \quad \text{a.s.}$$



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$$\sum_{k \in \mathbb{Z}} |\mathbb{L}_n(k)| e^{\beta k} < \infty \quad \text{a.s.}$$

The interval  $(\beta_-, \beta_+)$  may be infinite. For example, if for every  $n \in \mathbb{N}$  the support of a profile  $\{k \in \mathbb{Z} : \mathbb{L}_n(k) \neq 0\}$  is finite with probability one, then Assumption A1 holds with  $(\beta_-, \beta_+) = \mathbb{R}$ .



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Assume that there exist

- a sequence  $(w_n)_{n \in \mathbb{N}} \subset \mathbb{R}$  such that  $\lim_{n \rightarrow \infty} w_n = +\infty$ ;
- an open domain  $\mathcal{D} \subset \{\beta \in \mathbb{C} : \beta_- < \operatorname{Re} \beta < \beta_+\}$  such that  $\mathcal{D} \cap \mathbb{R} = (\beta_-, \beta_+)$ ;
- a deterministic analytic function  $\varphi : \mathcal{D} \rightarrow \mathbb{C}$  such that for every  $\beta \in (\beta_-, \beta_+)$  we have  $\varphi(\beta) \in \mathbb{R}$  and  $\varphi''(\beta) > 0$ .

From Assumption A1 it follows that with probability one, the normalized Laplace transform

$$W_n(\beta) := e^{-\varphi(\beta)w_n} \sum_{k \in \mathbb{Z}} \mathbb{L}_n(k) e^{\beta k}, \quad \beta \in \mathcal{D},$$

is a random analytic function on  $\mathcal{D}$  for every  $n \in \mathbb{N}$ .



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**Assumption A2:** With probability one the sequence of random analytic functions  $(W_n)_{n \in \mathbb{N}}$  converges locally uniformly on  $\mathcal{D}$ , as  $n \rightarrow \infty$ , to a random analytic function  $W_\infty$  such that  $\mathbb{P}[W(\beta) \neq 0 \text{ for all } \beta \in (\beta_-, \beta_+)] = 1$ .



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Assume also that the speed of convergence is superpolynomial in  $w_n$ :

**Assumption A3:** For every compact set  $K \subset \mathcal{D}$  and every  $r \in \mathbb{N}$  there exists an a.s. finite  $C_{K,r}$  such that for all  $n \in \mathbb{N}$ ,

$$\sup_{\beta \in K} |W_n(\beta) - W_\infty(\beta)| < C_{K,r} w_n^{-r}.$$



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The next assumption is technical. In the classical Edgeworth expansion for sums of i.i.d. variables it corresponds to the assumption that  $\mathbb{Z}$  is a minimal lattice containing the support of the step.





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**Assumption A4:** For every compact set  $K \subset (\beta_-, \beta_+)$ , every  $a > 0$  and  $r \in \mathbb{N}_0$ , we have

$$\sup_{\beta \in K} \left[ e^{-\varphi(\beta)w_n} \int_a^\pi \left| \sum_{k \in \mathbb{Z}} \mathbb{L}_n(k) e^{k(\beta+iu)} \right| du \right] = o(w_n^{-r}) \quad \text{a.s.}$$

as  $n \rightarrow \infty$ .



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The complete Bell polynomials  $B_j(z_1, \dots, z_j)$  are defined by the formal identity

$$\exp \left\{ \sum_{j=1}^{\infty} \frac{x^j}{j!} z_j \right\} = \sum_{j=0}^{\infty} \frac{x^j}{j!} B_j(z_1, \dots, z_j).$$

Therefore,  $B_0 = 1$  and for  $j \in \mathbb{N}$ ,

$$B_j(z_1, \dots, z_j) = \sum' \frac{j!}{i_1! \dots i_j!} \left( \frac{z_1}{1!} \right)^{i_1} \dots \left( \frac{z_j}{j!} \right)^{i_j},$$

where the sum  $\sum'$  is taken over all  $i_1, \dots, i_j \in \mathbb{N}_0$  such that  $1i_1 + 2i_2 + \dots + ji_j = j$ .



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The first few Bell polynomials:

$$B_1(z_1) = z_1, \quad B_2(z_1, z_2) = z_1^2 + z_2, \quad B_3(z_1, z_2, z_3) = z_1^3 + 3z_1z_2 + z_3.$$



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The “probabilist” Hermite polynomial  $\text{He}_n(x)$  of degree  $n$  is defined by

$$\text{He}_n(x) = e^{\frac{1}{2}x^2} \left( -\frac{d}{dx} \right)^n e^{-\frac{1}{2}x^2}.$$

The first few Hermite polynomials are:

$$\text{He}_1(x) = x, \quad \text{He}_2(x) = x^2 - 1, \quad \text{He}_3(x) = x^3 - 3x,$$

$$\text{He}_4(x) = x^4 - 6x^2 + 3, \quad \text{He}_6(x) = x^6 - 15x^4 + 45x^2 - 15.$$



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Consider a random profile  $\mathbb{L}_1, \mathbb{L}_2, \dots$  which satisfies assumptions A1–A4.

Consider the tilted profile  $k \mapsto e^{\beta k - \varphi(\beta)w_n} \mathbb{L}_n(k)$  and define its “mean”  $\mu(\beta)$  and the “standard deviation”  $\sigma(\beta)$ :

$$\mu(\beta) = \varphi'(\beta), \quad \sigma^2(\beta) = \varphi''(\beta),$$

and “the normalized coordinates”

$$x_n(k) = x_n(k; \beta) = \frac{k - \mu(\beta)w_n}{\sigma(\beta)\sqrt{w_n}}, \quad k \in \mathbb{Z}.$$



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Define the “deterministic cumulants”  $\kappa_j(\beta)$  and the “random cumulants”  $\chi_j(\beta)$  as follows:

$$\kappa_j(\beta) = \varphi^{(j)}(\beta), \quad \chi_j(\beta) = (\log W_\infty)^{(j)}(\beta).$$



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Let  $\mathbb{L}_1, \mathbb{L}_2, \dots$  be a random profile satisfying assumptions A1–A4. Fix  $r \in \mathbb{N}_0$  and a compact set  $K \subset (\beta_-, \beta_+)$ . Then

$$w_n^{\frac{r+1}{2}} \sup_{k \in \mathbb{Z}} \sup_{\beta \in K} \left| e^{\beta k - \varphi(\beta) w_n} \mathbb{L}_n(k) - \frac{W_\infty(\beta) e^{-\frac{1}{2} x_n^2(k)}}{\sigma(\beta) \sqrt{2\pi w_n}} \sum_{j=0}^r \frac{G_j(x_n(k); \beta)}{w_n^{j/2}} \right| \xrightarrow{\text{a.s.}} 0,$$

where  $G_j(x) = G_j(x; \beta)$ ,  $j \in \mathbb{N}_0$  is a polynomial of degree  $3j$ , defined by

$$G_j(x) = \frac{(-1)^j}{j!} e^{\frac{1}{2} x^2} B_j(D_1, \dots, D_j) e^{-\frac{1}{2} x^2}.$$

Here  $B_j$  is  $j$ -th Bell polynomial and  $D_1, D_2, \dots$  are linear differential operators with random coefficients:

$$D_j = \frac{\varphi^{(j+2)}(\beta)}{(j+1)(j+2)} \left( \frac{1}{\sigma(\beta)} \frac{d}{dx} \right)^{j+2} + \chi_j(\beta) \left( \frac{1}{\sigma(\beta)} \frac{d}{dx} \right)^j.$$



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The first three terms  $G_0, G_1, G_2$ :

$$G_0(x) = 1,$$

$$G_1(x) = \frac{\chi_1(\beta)}{\sigma(\beta)}x + \frac{\kappa_3(\beta)}{6\sigma^3(\beta)}\text{He}_3(x),$$

$$G_2(x) = \frac{\chi_1^2(\beta) + \chi_2(\beta)}{2\sigma^2(\beta)}\text{He}_2(x) + \left( \frac{\kappa_4(\beta)}{24\sigma^4(\beta)} + \frac{\kappa_3(\beta)\chi_1(\beta)}{6\sigma^4(\beta)} \right)\text{He}_4(x) \\ + \frac{\kappa_3^2(\beta)}{72\sigma^6(\beta)}\text{He}_6(x),$$





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Let  $Z_1, Z_2, \dots$  be an i.i.d. sequence of random variables with values in  $\mathbb{Z}$ , mean  $\mu := \mathbb{E}Z_1$ , variance  $\sigma^2 := \mathbb{D}Z_1 \neq 0$  and cumulant

$$\varphi(\beta) := \log \mathbb{E}e^{\beta Z_1},$$

which is finite on some interval  $(\beta_-, \beta_+)$  which contains 0.



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$$\varphi(\beta) := \log \mathbb{E}e^{\beta Z_1},$$

which is finite on some interval  $(\beta_-, \beta_+)$  which contains 0.

Consider a deterministic profile  $\mathbb{L}_n$ :

$$\mathbb{L}_n(k) = \mathbb{P}[Z_1 + \dots + Z_n = k], \quad k \in \mathbb{Z}.$$

Assumptions A1–A3 hold with  $\varphi$  defined above,  $w_n = n$  and  $W_\infty(\beta) = W_n(\beta) = 1$ . In particular, all cumulants  $\chi_k$  vanish. Assumption A4 holds if the distribution of  $Z_1$  is 1-arithmetic.



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From the main theorem with  $\beta = 0$  we obtain the classical Chebyshev-Edgeworth-Cramer expansion:

$$\lim_{n \rightarrow \infty} n^{\frac{r+1}{2}} \sup_{k \in \mathbb{N}} \left| \mathbb{P}[Z_1 + \dots + Z_n = k] - \frac{e^{-\frac{1}{2}x_n^2(k)}}{\sigma\sqrt{2\pi n}} \sum_{j=0}^r \frac{q_j(x_n(k))}{n^{j/2}} \right| = 0,$$

where  $q_j$  is a polynomial of degree  $3j$  with coefficients expressible in terms of the cumulants

$\kappa_2, \dots, \kappa_{j+2}$ .



# Branching random walks

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Consider a system of particles on  $\mathbb{Z}$  with the following evolution:

- at time 0 there is a single particle at 0;
- at each step *one* particle is chosen uniformly at random among existing particles;
- the chosen particle is replaced by a cluster of particles whose displacements w.r.t. the mother are described by a point process  $\zeta = \sum_{i=1}^N \delta_{Z_i}$  (where  $N$ , the number of particles, is a.s. finite) on  $\mathbb{Z}$ ;
- all random mechanisms involved in this definition are independent.



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Let  $S_n$  be the number of particles after  $n$  splitting events, and let their positions be  $x_{1,n}, \dots, x_{S_n,n}$ .

Let us denote by  $\mathbb{L}_n(k)$  the number of particles at site  $k \in \mathbb{Z}$  after  $n$  splitting events:

$$\mathbb{L}_n(k) = \#\{1 \leq j \leq S_n : x_{j,n} = k\}. \quad (1)$$

The function  $k \mapsto \mathbb{L}_n(k)$  is called the *profile* of the one-split BRW.



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Denote by  $\nu_k$  the expected number of particles at site  $k \in \mathbb{Z}$  in the cluster process  $\zeta$ :

$$\nu_k = \mathbb{E}\zeta(\{k\}) = \mathbb{E} \left[ \sum_{i=1}^N \mathbb{1}_{\{Z_i=k\}} \right], \quad k \in \mathbb{Z}.$$



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**Assumption B1:** We have  $\nu_k > 0$  for at least one  $k \in \mathbb{Z} \setminus \{0\}$ .



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**Assumption B1:** We have  $\nu_k > 0$  for at least one  $k \in \mathbb{Z} \setminus \{0\}$ .

**Assumption B2:** The cluster point process  $\zeta$  is a.s. non-empty, and the probability that it has at least 2 particles is positive. In other words,  $N \geq 1$  a.s. and  $\mathbb{P}[N = 1] \neq 1$ .





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Denote by  $m(\beta)$  the moment generating function of the intensity of the cluster point process  $\zeta$  minus 1:

$$m(\beta) = \sum_{k \in \mathbb{Z}} e^{\beta k} \nu_k - 1 = \mathbb{E} \left[ \sum_{i=1}^N e^{\beta Z_i} \right] - 1.$$

The expected number of particles at time  $n$  is  $\mathbb{E}S_n = 1 + m(0)n$ , where, by Assumption B2,

$$m(0) = \sum_{k \in \mathbb{Z}} \nu_k - 1 = \mathbb{E}N - 1 > 0.$$



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**Assumption B3:** The function  $m$  is finite on some non-empty open interval  $\mathcal{I}$  containing 0.



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**Assumption B3:** The function  $m$  is finite on some non-empty open interval  $\mathcal{I}$  containing 0.

**Assumption B4:**  $\nu$  is not concentrated on any proper additive subgroup of  $\mathbb{Z}$ . In other words,  $\nu(\mathbb{Z} \setminus a\mathbb{Z}) \neq 0$  for all  $a \in \{2, 3, \dots\}$ .



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Define the function

$$\varphi(\beta) = \frac{m(\beta)}{m(0)}, \quad \operatorname{Re} \beta \in \mathcal{I}$$

and denote by  $(\beta_-, \beta_+) \subset \mathcal{I}$  the open interval on which  $\varphi'(\beta)\beta < \varphi(\beta)$ :

$$\beta_- = \inf\{\beta \in \mathcal{I} : \varphi'(\beta)\beta < \varphi(\beta)\}, \quad (2)$$

$$\beta_+ = \sup\{\beta \in \mathcal{I} : \varphi'(\beta)\beta < \varphi(\beta)\}. \quad (3)$$

The interval  $(\beta_-, \beta_+)$  is non-empty because it contains 0.

**Assumption B5:** For any  $\beta \in (\beta_-, \beta_+)$  there is  $\gamma = \gamma(\beta) > 1$  such that

$$\mathbb{E} \left[ \left( \sum_{i=1}^N e^{\beta Z_i} \right)^\gamma \right] < \infty.$$



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The next theorem states that the sequence of the one-split BRW profiles satisfies Assumptions A2 and A3 with  $w_n = \log n$ . The proof uses an embedding into a continuous-time BRW in conjunction with well-known Biggins' results.



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## Theorem

*Under Assumptions B1–B3 and B5, there is an open neighborhood  $\mathcal{D}$  of the interval  $(\beta_-, \beta_+)$  in the complex plane such that, with probability 1,  $W_n$  converges to some random analytic function  $W_\infty$  locally uniformly on  $\mathcal{D}$ . Moreover, for every compact set  $K \subset \mathcal{D}$  and  $r \in \mathbb{N}$  we can find an a.s. finite random variable  $C_{K,r}$  such that for all  $n \in \mathbb{N}$ ,*

$$\sup_{\beta \in K} |W_n(\beta) - W_\infty(\beta)| < C_{K,r} (\log n)^{-r}.$$

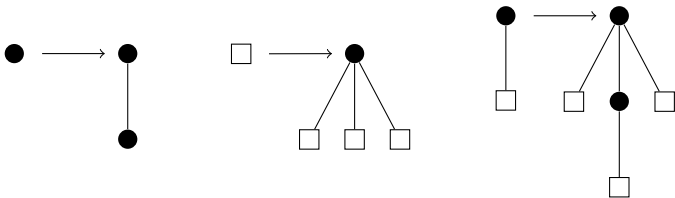
*With probability 1, the function  $W_\infty$  has no zeros on the interval  $(\beta_-, \beta_+)$ .*



# Random trees and one-split BRW

There are nodes of two types: the external ones (denoted by  $\square$ ) and the internal ones (denoted by  $\bullet$ ).

Construction rules for random trees:



Left: RRT. Middle:  $D$ -ary recursive with  $D = 3$ .  
Right: PORT.

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**Figure:** Sample realizations of random trees. Left: RRT. Middle:  $D$ -ary recursive tree with  $D = 3$ . Right: PORT.





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- particles correspond to (external or internal) nodes;
- positions of particles correspond to the depths of the nodes.



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Profile of a tree = the number of nodes at a given level

$\mathbb{L}_n(k)$  = the number of nodes at level  $k$  in a tree after  $n$  steps



# Random trees and the one-split BRW

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Mode and width

- particles correspond to (external or internal) nodes;
- positions of particles correspond to the depths of the nodes.

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B. Chauvin, L. Devroye, M. Drmota, M. Fuchs, H.-K. Hwang,  
H. Mahmoud, R. Neininger, A. Panholzer etc.



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(i) Binary search trees correspond to the one-split BRW with the deterministic displacement point process  $\zeta = 2\delta_1$  because at any step of the construction an external node at depth  $k$  is replaced by two new external nodes at depth  $k + 1$ . We have

$$\varphi(\beta) = 2e^\beta - 1, \quad m(0) = 1, \quad \mu(0) = \sigma^2(0) = \kappa_j(0) = 2, \quad j \in \mathbb{N}.$$

The constants  $\beta_- \approx -1.678$  and  $\beta_+ \approx 0.768$  are the solutions of  $2e^\beta(1 - \beta) = 1$ .



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(ii) Random recursive trees (RRTs) correspond to the one-split BRW with the deterministic displacement point process  $\zeta = \delta_0 + \delta_1$ . In particular,

$$\varphi(\beta) = e^\beta, \quad m(0) = 1, \quad \mu(0) = \sigma^2(0) = \kappa_j(0) = 1, \quad j \in \mathbb{N}.$$

We have  $\beta_- = -\infty$  and  $\beta_+ = 1$ .



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## Theorem

Let  $(\mathbb{L}_n(k))_{k \in \mathbb{Z}}$  be the profile of a binary search tree with  $n + 1$  external nodes. For every  $r \in \mathbb{N}_0$ , we have

$$\begin{aligned} & (\log n)^{\frac{r+1}{2}} \sup_{k \in \mathbb{Z}} \left| \frac{1}{n} \mathbb{L}_n(k) \right. \\ & \left. - \frac{e^{-\frac{(k-2 \log n)^2}{4 \log n}}}{\sqrt{4\pi \log n}} \sum_{j=0}^r G_j \left( \frac{k-2 \log n}{\sqrt{2 \log n}}; 0 \right) \frac{1}{(\log n)^{j/2}} \right| \xrightarrow{\text{a.s.}} 0, \end{aligned}$$

where  $G_j(x; 0)$  is a polynomial in  $x$  of degree  $3j$  whose coefficients can be linearly expressed through the derivatives  $W'_\infty(0), \dots, W_\infty^{(j)}(0)$ .



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The *mode*  $u_n$  and the *width*  $M_n$  of a profile are defined by

$$u_n = \arg \max_{k \in \mathbb{Z}} L_n(k), \quad M_n = \max_{k \in \mathbb{Z}} L_n(k).$$



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## Theorem

*There is an a.s. finite random variable  $K$  such that for  $n > K$ , the mode  $u_n$  of the BST with  $n + 1$  external nodes is equal to one of the numbers*

$$\lfloor 2 \log n + \chi_1(0) - 1/2 \rfloor \text{ or } \lceil 2 \log n + \chi_1(0) - 1/2 \rceil,$$

*where  $\lfloor \cdot \rfloor$ ,  $\lceil \cdot \rceil$  denote the floor and the ceiling functions, respectively, and  $\chi_1(0) = W'_\infty(0)/W_\infty(0)$ .*





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Actually, we can say more about the behavior of the mode. The following statements hold with probability 1:

- there are arbitrarily long intervals of consecutive  $n$ 's for which the mode  $u_n$  is unique and  $u_n = \lceil 2 \log n + \chi_1(0) - 1/2 \rceil$ ;
- similarly, there are arbitrarily long intervals of consecutive  $n$ 's for which  $u_n$  is unique and  $u_n = \lfloor 2 \log n + \chi_1(0) - 1/2 \rfloor$ ;
- the set of  $n \in \mathbb{N}$  such that  $u_n$  is the integer closest to  $2 \log n + \chi_1(0) - 1/2$  has asymptotic density one.



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## Theorem

Let  $M_n$  be the width of a binary search tree with  $n + 1$  external nodes. With probability 1, the set of subsequential limits of the sequence

$$\tilde{M}_n := 4 \log n \left( 1 - \frac{\sqrt{4\pi \log n} M_n}{n} \right), \quad n \in \mathbb{N},$$

is the interval

$$[\chi_2(0) - 1/12, \chi_2(0) + 1/6].$$

Set  $\theta_n = \min_{k \in \mathbb{Z}} |2 \log n + \chi_1(0) - 1/2 - k|$ . Then

$$\tilde{M}_n - \theta_n^2 \xrightarrow{a.s.} \chi_2(0) - \frac{1}{12}.$$



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# Thank you for attention!



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