

Convergence of complex Biggins martingale on the phase boundary

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Branching random walk

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We define an associated branching random walk:

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For a particle v , we write $|v|$ the generation in which v is born, $X(v)$ for its displacement from its parent, and for $S(v)$ for its position. Thus we have

$$S(v) = \sum_{u \leq v} X(u).$$

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Observe that

$$\begin{aligned}\mathbb{E}[Z_{n+1}|\mathcal{F}_n] &= \mathbb{E}\left[\sum_{|v|=n+1} e^{-\lambda S(v)}|\mathcal{F}_n\right] \\ &= \mathbb{E}\left[\sum_{|v|=n} \sum_{v < u, |u|=n+1} e^{-\lambda(S(v)+X(u))}|\mathcal{F}_n\right] = Z_n \mathbb{E}[Z_1].\end{aligned}$$

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In particular, for $m(\lambda) = \mathbb{E} \sum_i e^{-\lambda X_i}$, the process

$$W_n(\lambda) := \frac{Z_n(\lambda)}{\mathbb{E}Z_n(\lambda)} = m^{-n}(\lambda)\mathbb{E}[Z_n(\lambda)]$$

is a martingale, called Biggins' martingale.

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- ▶ When the limit is nontrivial?
- ▶ What is the rate of convergence?

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- ▶ $W_n(\lambda)$ is a “toy model” for the Gaussian multiplicative chaos i.e. measures of the form

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where h is log-correlated Gaussian field, i.e.

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- ▶ applications in physics

Real case

Let assume that $d = 1$ and $\theta \in \mathbb{R}$. By martingale convergence theorem $W_n(\theta) \rightarrow W(\theta)$ for some $W(\theta) \geq 0$.

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Theorem (Biggins 1977)

The following are equivalent

- ▶ $\mathbb{P}[W(\theta) = 0] = \mathbb{P}[\text{the branching process gets extinct}]$
- ▶ $\mathbb{P}[W(\theta) = 0] < 1$
- ▶ $\mathbb{E}W(\theta) = 1$
- ▶ $\mathbb{E}[Z_1(\theta) \log^+ Z_1(\theta)] < \infty$ and $\theta m'(\theta)/m(\theta) < \log m(\theta)$

Complex case

It is convenient to replace $e^{-\lambda S(v)}/m(\lambda)^n$ by a random variable $L(v)$.

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$$\{L(v)^{-1}L(u)\}_{u:u'=v},$$

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We define the moment generating function φ by

$$\varphi(t) = \mathbb{E} \sum_{|v|=1} |L(v)|^t$$

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Theorem (Biggins 1992)

Suppose that $\varphi(p) < 1$ for some $p \in (1, 2]$, then W_n converges almost surely and in L^p to some limit W .

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Theorem (K, Meiners)

Suppose that $\varphi(\alpha) = 1$ and $\varphi'(\alpha) = 0$ for some $\alpha \in (1, 2)$ then W_n converges almost surely and in L^p for any $p < \alpha$.

If $\varphi(2) = 1$ then W_n does not converge in probability.

Gaussian splitting

Let $\mathcal{L} = \{X_1, X_2\}$ where X_i are independent standard normal variables. In this case $m(\lambda) = 2e^{\lambda^2/2}$ and $\phi(t) = 2^{1-t} e^{\frac{t}{2}((t+1)\theta^2 - \eta^2)}$

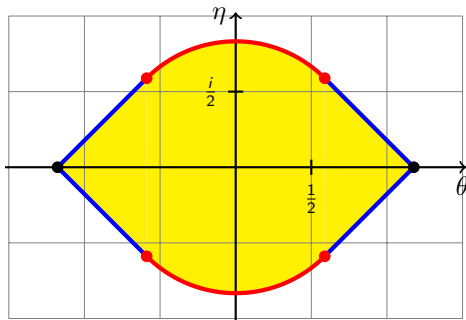


Figure: In the above example the yellow part corresponds to those $\lambda = \theta + i\eta$ for which $\phi(p) < 1$ for some $p \in (1, 2]$. The blue lines corresponds to the case $\phi(\alpha) = 1$ and $\phi'(\alpha) = 0$ for some $\alpha \in (1, 2)$. The red curve corresponds to the case $\phi(2) = 1$.

Change of the measure

Suppose that

$$\mathbb{E} \sum_{|v|=1} |L(v)|^\alpha = 1.$$

Then we can define a random variable Y by

$$\mathbb{E}[f(Y)] = \mathbb{E} \sum_{|v|=1} f(\log |L(v)|) |L(v)|^\alpha,$$

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Many to one lemma

For any function $F : \mathbb{R}^n \mapsto \mathbb{R}$ we have

$$\mathbb{E} \left[\sum_{|v|=n} F(S(v_0), S(v_1), \dots, S(v_n)) |L(v)|^\alpha \right] = \mathbb{E} F(S_0, S_1, \dots, S_n),$$

where v_i is an ancestor of v from the i -th generation.

Idea of the proof: Naive approach

First we can try to estimate the second moment of the martingale:

$$\begin{aligned}\mathbb{E}|W_n - 1|^2 &= \sum_{i \leq n-1} \mathbb{E}|W_{i+1} - W_i|^2 \\ &= \sum_{i \leq n-1} \mathbb{E} \sum_{|v|=i} |L(v)|^2 \left| \sum_{u:u'=v} L(v)^{-1} L(u) - 1 \right|^2 \\ &= \mathbb{E} \left| \sum_{|u|=1} L(u) - 1 \right|^2 \mathbb{E} \sum_{|v| \leq n-1} |L(v)|^2 \\ &= C(1 + \phi(2) + \dots + \phi^{n-1}(2)),\end{aligned}$$

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Since $\phi(2) > 1$ the above sum diverges.

Reason: $\max_v |L(v)|$ has too heavy tail.

Idea of the proof: estimation of the maximum

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We have

$$\begin{aligned}\mathbb{P}[\max_v L(v) > t] &\leq \mathbb{E}[\#\{v : L(v) > t, L(u) \leq t \text{ for all } u \leq v\}] \\ &= \sum_v \mathbb{E} \mathbf{1}_{[L(v) > t, L(u) \leq t \text{ for all } u \leq v]} \\ &= \sum_n \mathbb{E} e^{-\alpha S_n} \mathbf{1}_{[S_n > \log t, S_k \leq t \text{ for all } k \leq n]} \leq t^{-\alpha}\end{aligned}$$

Idea of the proof: truncation argument

For given t , we define a new martingale $\{W_n^t\}$ by $W_0^t = 1$ and

$$W_{i+1}^t - W_i^t = \sum_{|v|=i} L(v) \mathbf{1}_{[|L(w)| \leq t \text{ for all } w \leq v]} \left(\sum_{u: u' = v} L(v)^{-1} L(u) - 1 \right).$$

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We can compute the second moment:

$$\begin{aligned} \mathbb{E}|W_n^t - 1|^2 &= \sum_{i \leq n-1} \mathbb{E}|W_{i+1}^t - W_i^t|^2 \\ &= C \sum_{i \leq n-1} \mathbb{E}|L(v)|^2 \mathbf{1}_{[|L(w)| \leq t \text{ for all } w \leq v]} \\ &= C \sum_{i \leq n-1} \mathbb{E} e^{(2-\alpha)S_i} \mathbf{1}_{[S_j \leq \log t \text{ for all } j \leq i]} \lesssim t^{2-\alpha} \end{aligned}$$

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In particular, W_n is L^p bounded martingale for any $p < \alpha$. □

Rate of convergence (work in progress)

We have the following tail behavior

Theorem (Iksanov, K, Meiners)

For the limit W of the martingale W_n we have

$$t^\alpha \mathbb{P}[|W| > t] \rightarrow K > 0,$$

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As a consequence we obtain

Theorem (Iksanov, K, Meiners)

$$n^{\frac{1}{2\alpha}} (W - W_n) \xrightarrow{d} \sqrt[\alpha]{D_\infty} S_\alpha,$$

where D_∞ is the limit of derivative martingale and S_α is independent α stable distribution.