

THE MONOTONE POISSON PROCESS

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Abstract. The coefficients of the moments of the monotone Poisson law are shown to be a type of Stirling number of the first kind; certain combinatorial identities relating to these numbers are proved and a new derivation of the Cauchy transform of this law is given. An investigation is begun into the classical Azéma-type martingale which corresponds to the compensated monotone Poisson process; it is shown to have the chaotic-representation property and its sample paths are described.

1. Introduction. If X is a random variable distributed according to the monotone Poisson law, i.e., Muraki's monotonic law of small numbers [16, Section 4], with mean $t > 0$ then [4, Section 4]

$$(1) \quad \mathbb{E}[X^{n+1}] = t + \sum_{m=1}^n \left(\sum_{2 \leq j_1 < \dots < j_m \leq n+1} j_1 \cdots j_m \right) \frac{t^{m+1}}{(m+1)!} \quad \forall n \geq 1.$$

The coefficients appearing in (1) are shown below to be a type of generalized Stirling number. Using a method of proof from our earlier work, closed forms for exponential generating functions of such Stirling numbers are derived; it seems these have appeared previously in the literature only for sporadic examples [19]. The Cauchy transform for the monotone Poisson law, found by Muraki in [16], is derived in a novel manner from the recurrence relation satisfied by the moment coefficients.

In [4] the monotone Poisson process was realised as a vacuum-adapted semimartingale; the corresponding regular quantum semimartingale, i.e., its HP-adapted projection [3], is used to construct a classical semimartingale $y = (y_t)_{t \geq 0}$ distributed according to the monotone Poisson law. It is shown that $y_t = x_t + t$ for all $t \geq 0$, where x is a normal martingale (the *compensated monotone Poisson process*) which satisfies the structure equation

$$(2) \quad x_0 = 0, \quad d[x]_t = (1 - t - x_{t-}) dx_t + dt \quad \forall t \geq 0.$$

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This martingale x is shown to have the chaotic-representation property and to be purely discontinuous, with sample paths as follows: if $g_t := \sup\{s \in [0, t] : x_s = 1 - s\}$ is the last time up to t that $x_s = 1 - s$ then

$$x_t = -W_\bullet(-\exp(-1 - t + g_t)) - t \quad \forall t \geq 0,$$

where W_\bullet is one of the two real branches of the Lambert W function (i.e., the multi-valued inverse to the function $z \mapsto ze^z$). A numerical simulation is employed to produce approximations to these paths.

1.1. Notation. The set of natural numbers is denoted by \mathbb{N} and equals $\{1, 2, 3, \dots\}$; the set of non-negative integers is denoted by \mathbb{Z}_+ and equals $\{0, 1, 2, \dots\}$. If $\{f_\alpha : \alpha \in A\}$ is a collection of real-valued functions with common domain Ω then $\sigma(f_\alpha : \alpha \in A)$ is the smallest σ -algebra on Ω to make each function measurable. If A is a subset of some algebra then the subalgebra that it generates is denoted by $\text{alg } A$.

1.2. Conventions. An empty sum equals 0; an empty product equals 1. A function $f : A \rightarrow B$ has domain $\text{Dom } f = A$ and range $\text{Ran } f = f(A) \subseteq B$. The binomial coefficient $\binom{n}{m}$ equals 0 if $m < 0$.

2. A recurrence relation

DEFINITION 1. For all $a \in \mathbb{Z}_+$ define

$$\left| \begin{array}{c} n \\ a \end{array} \right| : \mathbb{N} \times \mathbb{N} \rightarrow \mathbb{Z}_+; (n, m) \mapsto \left| \begin{array}{c} n \\ m \\ a \end{array} \right|$$

by setting $\left| \begin{array}{c} n \\ 1 \\ a \end{array} \right| = 1$ for all $n \in \mathbb{N}$, $\left| \begin{array}{c} n \\ m \\ a \end{array} \right| = 0$ if $m > n$ and by using the recurrence relation

$$(3) \quad \left| \begin{array}{c} n+1 \\ m+1 \\ a \end{array} \right| = \left| \begin{array}{c} n \\ m+1 \\ a \end{array} \right| + (n-1+a) \left| \begin{array}{c} n \\ m \\ a \end{array} \right| \quad \forall n, m \in \mathbb{N}.$$

These numbers are *generalized Stirling numbers of the first kind*; this terminology is justified by Remark 5 and Remark 6 below.

PROPOSITION 2. The function $\left| \begin{array}{c} n \\ a \end{array} \right|$ is well defined for all $a \in \mathbb{Z}_+$ and satisfies

$$\left| \begin{array}{c} n \\ n \\ a \end{array} \right| = \prod_{j=0}^{n-2} (j+a) = \begin{cases} \delta_1^n & \text{if } a = 0, \\ (n-2+a)!/(a-1)! & \text{if } a \in \mathbb{N}, \end{cases}$$

for all $n \in \mathbb{N}$, where δ is the Kronecker delta ($\delta_b^a = 1$ if $a = b$ and 0 otherwise).

Proof. Induction. ■

The recurrence relation (3) and the boundary conditions provide a quick method for the computation of $\left| \begin{array}{c} n \\ a \end{array} \right|$; a range of values for $\left| \begin{array}{c} n \\ 2 \end{array} \right|$ is given in Table 1.

PROPOSITION 3.

$$\left| \begin{array}{c} n+1 \\ m+1 \\ a \end{array} \right| = \sum_{a \leq j_1 < \dots < j_m \leq n-1+a} j_1 \cdots j_m$$

for all $n, m \in \mathbb{N}$, where the summation is taken over integral j_1, \dots, j_m .

$n \backslash m$	1	2	3	4	5	6	7	8	9	10
1	1	0	0	0	0	0	0	0	0	0
2	1	2	0	0	0	0	0	0	0	0
3	1	5	6	0	0	0	0	0	0	0
4	1	9	26	24	0	0	0	0	0	0
5	1	14	71	154	120	0	0	0	0	0
6	1	20	155	580	1044	720	0	0	0	0
7	1	27	295	1665	5104	8028	5040	0	0	0
8	1	35	511	4025	18424	48860	69264	40320	0	0
9	1	44	826	8624	54649	214676	509004	663696	362880	0
10	1	54	1266	16884	140889	761166	2655764	5753736	6999840	3628800

Table 1. $\left| \begin{smallmatrix} n \\ m \end{smallmatrix} \right|_2$ for $n, m = 1, \dots, 10$.

Proof. Define $f_a: \mathbb{N} \times \mathbb{N} \rightarrow \mathbb{Z}_+$ by setting $f_a(n, 1) = 1$ for all $n \in \mathbb{N}$ and

$$f_a(n + 1, m + 1) = \sum_{a \leq j_1 < \dots < j_m \leq n-1+a} j_1 \cdots j_m \quad \forall n, m \in \mathbb{N};$$

note that $f_a(n, m) = 0$ if $m > n$. Furthermore, if $n, m \in \mathbb{N}$ are such that $2 \leq m \leq n$ then

$$\begin{aligned} f_a(n + 1, m + 1) &= \sum_{a \leq j_1 < \dots < j_m \leq n-2+a} j_1 \cdots j_m + \sum_{a \leq j_1 < \dots < j_{m-1} \leq n-2+a} j_1 \cdots j_{m-1} (n - 1 + a) \\ &= f_a(n, m + 1) + (n - 1 + a) f_a(n, m); \end{aligned}$$

this identity also holds if $m > n$ or $m = 1$:

$$f_a(n + 1, m + 1) = 0 = 0 + 0 = f_a(n, m + 1) + f_a(n, m) \quad \forall m > n$$

and

$$f_a(n + 1, 2) = \sum_{j_1=a}^{n-1+a} j_1 = f_a(n, 2) + (n - 1 + a) = f_a(n, 2) + (n - 1 + a) f_a(n, 1).$$

Thus f_a and $\left| \begin{smallmatrix} \cdot \\ a \end{smallmatrix} \right|_a$ both satisfy the recurrence relation (3), with the same boundary conditions; the result follows. ■

COROLLARY 4. *If X is a random variable with the monotone Poisson distribution and mean $t > 0$ then*

$$\mathbb{E}[X^n] = \sum_{m=1}^{\infty} \left| \begin{smallmatrix} n \\ m \end{smallmatrix} \right|_2 \frac{t^m}{m!} \quad \forall n \in \mathbb{N}.$$

Proof. This is immediate from Proposition 3 and (1). ■

The definition of $\left| \begin{smallmatrix} \cdot \\ 2 \end{smallmatrix} \right|_2$ extends to $\mathbb{Z}_+ \times \mathbb{Z}_+$ by setting $\left| \begin{smallmatrix} 0 \\ 0 \end{smallmatrix} \right|_2 = 1$ and $\left| \begin{smallmatrix} n \\ 0 \end{smallmatrix} \right|_2 = \left| \begin{smallmatrix} 0 \\ n \end{smallmatrix} \right|_2 = 0$ for all $n \in \mathbb{N}$; the identity (3) holds for all $n, m \in \mathbb{Z}_+$ and, in the notation of Corollary 4,

$$\mathbb{E}[X^n] = \sum_{m=0}^{\infty} \left| \begin{smallmatrix} n \\ m \end{smallmatrix} \right|_2 \frac{t^m}{m!} \quad \forall n \in \mathbb{Z}_+.$$

REMARK 5. By Proposition 3, if $n \in \mathbb{N}$ then

$$\prod_{j=0}^{n-1} (x + j + a) = x^n + \sum_{m=1}^n \left(\sum_{a \leq j_1 < \dots < j_m \leq n-1+a} j_1 \cdots j_m \right) x^{n-m} = \sum_{m=0}^n \left| \begin{matrix} n+1 \\ m+1 \end{matrix} \right|_a x^{n-m}.$$

Hence, in the notation of Mitrinović [15] (who appears to have been the first to study such numbers),

$$\left| \begin{matrix} n+1 \\ m+1 \end{matrix} \right|_a = R_n^{n-m}(-a, -1) = (-1)^{n-m} R_n^{n-m}(a, 1) \quad \forall n \in \mathbb{N}, m = 0, \dots, n.$$

In terms of Broder’s r -Stirling numbers of the first kind [5],

$$\left| \begin{matrix} n+1 \\ m+1 \end{matrix} \right|_a = \left[\begin{matrix} n+a \\ n+a-m \end{matrix} \right]_a \quad \forall n, m \in \mathbb{Z}_+$$

and, in the unified approach of Hsu and Shiue [12],

$$\left| \begin{matrix} n+1 \\ m+1 \end{matrix} \right|_a = S^1(n, n-m; -1, 0, a) = R_1(n, n-m, a) \quad \forall n \in \mathbb{Z}_+, m = 0, \dots, n,$$

where $R_1(n, k, \lambda)$ is the r -Stirling polynomial of the first kind due to Carlitz [6, 7].

REMARK 6. If $\left[\begin{matrix} n \\ m \end{matrix} \right]$ denotes the unsigned Stirling numbers of the first kind, following Karamata *via* Graham, Knuth and Patashnik [10], then

$$\left[\begin{matrix} n \\ m \end{matrix} \right] = \left| \begin{matrix} n+1 \\ n+1-m \end{matrix} \right|_0 \quad \forall n \in \mathbb{Z}_+, m \in \mathbb{Z} \cap]-\infty, n].$$

Gupta [11, Chapter 5] introduces numbers $G(n, m)$ such that $G(n, m) = \left| \begin{matrix} n+1 \\ m+1 \end{matrix} \right|_1$ for all $n, m \in \mathbb{Z}_+$. Since

$$\sum_{m=0}^{n+1} \left| \begin{matrix} n+2 \\ m+1 \end{matrix} \right|_0 x^{n+1-m} = \prod_{j=0}^n (x + j) = x \prod_{j=0}^{n-1} (x + j + 1) = \sum_{m=0}^n \left| \begin{matrix} n+1 \\ m+1 \end{matrix} \right|_1 x^{n+1-m} \quad \forall n \in \mathbb{N},$$

$\left| \begin{matrix} n+2 \\ m+1 \end{matrix} \right|_0 = \left| \begin{matrix} n+1 \\ m+1 \end{matrix} \right|_1$ for all $n \in \mathbb{N}$ and all $m \in \mathbb{Z}_+$ with $m \leq n$. It is readily verified that this identity holds for all $n, m \in \mathbb{Z}_+$ and

$$G(n, m) = \left| \begin{matrix} n+1 \\ m+1 \end{matrix} \right|_1 = \left| \begin{matrix} n+2 \\ m+1 \end{matrix} \right|_0 = \left[\begin{matrix} n+1 \\ n+1-m \end{matrix} \right],$$

i.e., Gupta’s numbers are unsigned Stirling numbers of the first kind.

A result of Mitrinović [15, (1.4)] implies that

$$\left| \begin{matrix} n \\ m \end{matrix} \right|_a = \sum_{j=0}^{m-1} \left| \begin{matrix} n \\ m-j \end{matrix} \right|_0 \binom{n-m+j}{n-m} a^j \quad \forall a \in \mathbb{Z}_+.$$

The following transformation theorem is a generalization of this result.

THEOREM 7. If $a, b \in \mathbb{Z}_+$ then

$$\left| \begin{matrix} n \\ m \end{matrix} \right|_a = \sum_{j=0}^{m-1} \left| \begin{matrix} n \\ m-j \end{matrix} \right|_b \binom{n-m+j}{n-m} (a-b)^j \quad \forall n, m \in \mathbb{N}.$$

Proof. If $n \in \mathbb{N}$ then, by Remark 5,

$$\prod_{j=0}^{n-1} (x + j + b) = \sum_{m=0}^n \left| \begin{matrix} n+1 \\ m+1 \end{matrix} \right|_b x^{n-m} = \sum_{l=0}^n \left| \begin{matrix} n+1 \\ n+1-l \end{matrix} \right|_b x^l$$

and so, since $x + j + a = (x + a - b) + j + b$,

$$\begin{aligned} \prod_{j=0}^{n-1} (x + j + a) &= \sum_{l=0}^n \left| \begin{matrix} n+1 \\ n+1-l \end{matrix} \right|_b (x + a - b)^l \\ &= \sum_{l=0}^n \left| \begin{matrix} n+1 \\ n+1-l \end{matrix} \right|_b \sum_{m=0}^l \binom{l}{m} (a-b)^{l-m} x^m \\ &= \sum_{m=0}^n \sum_{l=m}^n \left| \begin{matrix} n+1 \\ n+1-l \end{matrix} \right|_b \binom{l}{m} (a-b)^{l-m} x^m. \end{aligned}$$

Hence, if $m \in \mathbb{Z}_+$ is such that $0 \leq m \leq n$,

$$\begin{aligned} \left| \begin{matrix} n+1 \\ n+1-m \end{matrix} \right|_a &= \sum_{l=m}^n \left| \begin{matrix} n+1 \\ n+1-l \end{matrix} \right|_b \binom{l}{m} (a-b)^{l-m} \\ &= \sum_{j=0}^{n-m} \left| \begin{matrix} n+1 \\ n+1-m-j \end{matrix} \right|_b \binom{m+j}{m} (a-b)^j \end{aligned}$$

and the result follows. (The cases $m > n$ and $m = n = 1$ are verified immediately upon inspection.) ■

REMARK 8. Theorem 7, with $b = 0$, and Remark 6 yield the identities

$$\left| \begin{matrix} n+k+1 \\ n+1 \end{matrix} \right|_a = \sum_{j=0}^n \left[\begin{matrix} n+k \\ j+k \end{matrix} \right] \binom{j+k}{k} a^j \quad \forall n, k, a \in \mathbb{Z}_+,$$

which correspond, in the cases $a = 2, \dots, 4$, $k = 1, \dots, 4$ (and $k = 5$ if $a = 2$) to formulae [19, A001705–A001709, A001711–A001714, A001716–A001719, A001721–A001724] given by Crstici.

3. Generating functions. Recall that the *exponential generating function* for a sequence $(b_n)_{n \geq 0}$ is the function determined by the power series $\sum_{n=0}^{\infty} b_n z^n / n!$.

DEFINITION 9. For all $a \in \mathbb{Z}_+$ define

$$g_{a,n}(w) := \sum_{m=1}^{\infty} \left| \begin{matrix} n \\ m \end{matrix} \right|_a \frac{w^m}{m!} = \sum_{m=1}^n \left| \begin{matrix} n \\ m \end{matrix} \right|_a \frac{w^m}{m!} \quad \forall n \in \mathbb{N}, w \in \mathbb{C}.$$

Note that $g_{a,n}(0) = 0$, $g_{a,1}(w) = w$ and

$$g_{a,n+1}(w) = w + \sum_{m=2}^{\infty} \left| \begin{matrix} n+1 \\ m \end{matrix} \right|_a \frac{w^m}{m!} = w + \sum_{m=2}^{\infty} \left| \begin{matrix} n \\ m \end{matrix} \right|_a \frac{w^m}{m!} + (n-1+a) \sum_{m=2}^{\infty} \left| \begin{matrix} n \\ m-1 \end{matrix} \right|_a \frac{w^m}{m!},$$

so

$$(4) \quad (g_{a,n+1} - g_{a,n})' = (n-1+a)g_{a,n} \quad \forall n \in \mathbb{N}.$$

THEOREM 10. *If $a \in \mathbb{N}$ and*

$$(5) \quad h_{a,n}(w, z) := \frac{1}{n!} \left(\frac{\partial}{\partial z} \right)^{n-1} \frac{(w - \log(1-z))^n}{(1-z)^a} \quad (w \in \mathbb{C}, |z| < 1)$$

for all $n \in \mathbb{N}$ (where \log denotes the principal branch of the natural logarithm) then $g_{a,n}(w) = h_{a,n}(w, 0)$ for all $n \in \mathbb{N}$ and all $w \in \mathbb{C}$.

Proof. Note first that $h_{a,1}(w, 0) = w$ and [10, (5.56)]

$$(6) \quad \frac{1}{(1-z)^a} = \sum_{l=0}^{\infty} \binom{l+a-1}{a-1} z^l \quad (|z| < 1).$$

Next, letting $\partial_z^m := \frac{\partial^m}{\partial z^m}$ et cetera,

$$\begin{aligned} nh_{a,n}(w, z) &= \frac{1}{(n-1)!} \sum_{m=0}^{n-1} \binom{n-1}{m} \partial_z^m (w - \log(1-z))^n \partial_z^{n-1-m} \sum_{l=0}^{\infty} \binom{l+a-1}{a-1} z^l \\ &= \sum_{m=0}^{n-1} \frac{1}{m!} j_{m,n}(w, z) \left(\binom{n-m+a-2}{a-1} + (n-m) \binom{n-m+a-1}{a-1} z + \dots \right) \\ &= \sum_{m=0}^{n-1} \frac{(n-m+a-2)!}{m!(n-m-1)!} j_{m,n}(w, z) (1 + (n-m+a-1)z + \dots), \end{aligned}$$

where $j_{m,n}(w, z) = \partial_z^m (w - \log(1-z))^n$ and the ellipses denote a term of order z^2 . If

$$r_{a,n} := \partial_w (h_{a,n+1} - h_{a,n}) - (n-1+a)h_{a,n} = (\partial_z - \partial_w - (n-1+a))h_{a,n}$$

then

$$nr_{a,n}(w, 0) = \sum_{m=0}^{n-1} \frac{(n-m-2+a)!}{m!(n-m-1)!} (j_{m+1,n}(w, 0) - nj_{m,n-1}(w, 0) - mj_{m,n}(w, 0))$$

but $j_{m+1,n}(w, 0) = mj_{m,n}(w, 0) + nj_{m,n-1}(w, 0)$ [4, Proof of Proposition 4.1] and so $r_{a,n}(w, 0) = 0$ for all $w \in \mathbb{C}$ and all $n \in \mathbb{N}$. Thus the differential equation (4) is satisfied, with $g_{a,n}$ replaced by $w \mapsto h_{a,n}(w, 0)$ et cetera; as $h_{a,n}(0, 0) = 0$ for all $n \in \mathbb{N}$, because $j_{m,n}(0, 0) = 0$ if $0 \leq m < n$ [4, Proof of Proposition 4.1], the result follows. ■

LEMMA 11. *Given any $a, k \in \mathbb{Z}_+$ not both zero, the power series*

$$(7) \quad \sum_{n=1}^{\infty} \binom{n+k}{n} \Big|_a \frac{z^n}{n!}$$

has radius of convergence 1. If $a = k = 0$, the series (7) equals z .

Proof. If $a \in \mathbb{N}$ then, by Proposition 3,

$$\binom{n+k}{n} \prod_{j=a}^{n-1+a} j \leq \binom{n+1+k}{n+1} \Big|_a = \sum_{a \leq j_1 < \dots < j_n \leq n+k-1+a} j_1 \cdots j_n \leq \binom{n+k}{n} \prod_{j=k+a}^{n+k-1+a} j$$

for all $k \in \mathbb{Z}_+$ and all $n \in \mathbb{N}$, so if

$$b_n := \frac{1}{n!} \binom{n-1+k}{n-1} \prod_{j=a}^{n-2+a} j = \binom{n-1+k}{n-1} \frac{(n-2+a)!}{n!(a-1)!}$$

and

$$c_n := \frac{1}{n!} \binom{n-1+k}{n-1} \prod_{j=k+a}^{n+k-2+a} j = \binom{n-1+k}{n-1} \frac{(n+k-2+a)!}{n!(k-1+a)!}$$

then $b_n \leq \frac{1}{n!} \left| \binom{n+k}{n} \right| \leq c_n$ for all $n \in \mathbb{N}$; the result follows by the ratio and comparison tests. If $a = 0$ then, since $\left| \binom{n+2}{m+1} \right|_0 = \left| \binom{n+1}{m+1} \right|_1$ for all $n, m \in \mathbb{Z}_+$ (Remark 6),

$$\sum_{n=1}^{\infty} \left| \binom{n+k}{n} \right|_0 \frac{z^n}{n!} = \sum_{n=1}^{\infty} \left| \binom{n+k-1}{n} \right|_1 \frac{z^n}{n!}$$

has radius of convergence 1 for all $k \in \mathbb{N}$; the final case follows from Proposition 2. ■

The next result was proved by Broder [5, Theorem 15] using combinatorial methods.

COROLLARY 12. *Given any $a, k \in \mathbb{N}$, the exponential generating function for the sequence $(b_n)_{n \geq 0}$, where $b_0 = 0$ and $b_n = \left| \binom{n+k}{n} \right|_a$ for all $n \in \mathbb{N}$, equals*

$$\frac{1}{k!} \left(\frac{d}{dz} \right)^{k-1} \frac{(-\log(1-z))^k}{(1-z)^a} \quad (|z| < 1).$$

Proof. If $n, k \in \mathbb{N}$ then Theorem 10 implies that

$$\begin{aligned} \left| \binom{n+k}{n} \right|_a &= \left(\frac{\partial}{\partial w} \right)^n \frac{1}{(n+k)!} \left(\frac{\partial}{\partial z} \right)^{n+k-1} \frac{(w - \log(1-z))^{n+k}}{(1-z)^a} \Big|_{z=0, w=0} \\ &= \frac{1}{(n+k)!} \left(\frac{d}{dz} \right)^{n+k-1} \frac{(n+k)!}{k!} \frac{(-\log(1-z))^k}{(1-z)^a} \Big|_{z=0} \\ &= \left(\frac{d}{dz} \right)^n \left(\frac{1}{k!} \left(\frac{d}{dz} \right)^{k-1} \frac{(-\log(1-z))^k}{(1-z)^a} \right) \Big|_{z=0}, \end{aligned}$$

as required. ■

Amusingly, Corollary 12 is also true for $k = 0$, if $(d/dz)^{-1}$ is interpreted as integration. More precisely, if $a \in \mathbb{N}$ then, by Proposition 2,

$$\sum_{n=1}^{\infty} \left| \binom{n}{n} \right|_a \frac{z^n}{n!} = \sum_{m=0}^{\infty} \binom{m+a-1}{a-1} \int_0^z w^m dw = \int_0^z \frac{1}{(1-w)^a} dw,$$

where the second equality follows from (6). Hence

$$(8) \quad \sum_{n=1}^{\infty} \left| \binom{n}{n} \right|_a \frac{z^n}{n!} = \begin{cases} z & \text{if } a = 0, \\ -\log(1-z) & \text{if } a = 1, \\ \frac{1}{a-1} \left(\frac{1}{(1-z)^{a-1}} - 1 \right) & \text{if } a \geq 2. \end{cases}$$

REMARK 13. By Remark 6, Corollary 12 and (8), if $k \geq 2$ then

$$\sum_{n=1}^{\infty} \left| \binom{n+k}{n} \right|_0 \frac{z^n}{n!} = \sum_{n=1}^{\infty} \left| \binom{n+k-1}{n} \right|_1 \frac{z^n}{n!} = \frac{1}{(k-1)!} \left(\frac{d}{dz} \right)^{k-2} \frac{(-\log(1-z))^{k-1}}{1-z},$$

$$\sum_{n=1}^{\infty} \left| \begin{matrix} n+1 \\ n \end{matrix} \right|_0 \frac{z^n}{n!} = \sum_{n=1}^{\infty} \left| \begin{matrix} n \\ n \end{matrix} \right|_1 \frac{z^n}{n!} = -\log(1-z) \quad \text{and} \quad \sum_{n=1}^{\infty} \left| \begin{matrix} n \\ n \end{matrix} \right|_0 \frac{z^n}{n!} = z.$$

If $k \in \mathbb{Z}_+$ then, by Remark 6,

$$\sum_{n=1}^{\infty} \left| \begin{matrix} n+k \\ n \end{matrix} \right|_0 \frac{z^n}{n!} = \sum_{n=1}^{\infty} \left[\begin{matrix} n+k-1 \\ k \end{matrix} \right] \frac{z^n}{n!} = \int_0^z \sum_{m=0}^{\infty} \left[\begin{matrix} m+k \\ k \end{matrix} \right] \frac{w^m}{m!} dw$$

and so [10, (7.50)]

$$\sum_{n=0}^{\infty} \left[\begin{matrix} n+k \\ k \end{matrix} \right] \frac{z^n}{n!} = \frac{1}{k!} \left(\frac{d}{dz} \right)^k (-\log(1-z))^k \quad \forall k \in \mathbb{Z}_+.$$

4. A Cauchy transform. If X is a random variable distributed according to the measure μ then the *Cauchy transform* (or *Cauchy-Stieltjes transform*) of X is

$$G_X : \mathbb{C} \setminus \text{supp } \mu \rightarrow \mathbb{C}; \quad z \mapsto \int_{\text{supp } \mu} \frac{d\mu(x)}{z-x}.$$

This function is holomorphic on $\mathbb{C} \setminus \text{supp } \mu$ and, if $\text{supp } \mu \subseteq [a, b]$,

$$G_X(z) = \frac{1}{z} \int_a^b \sum_{n=0}^{\infty} \frac{x^n}{z^n} d\mu(x) = \sum_{n=0}^{\infty} \mathbb{E}[X^n] z^{-n-1} \quad (|z| > \max\{|a|, |b|\}).$$

LEMMA 14. *If $m \in \mathbb{N}$ then the series*

$$\sum_{n=1}^{\infty} \left| \begin{matrix} n \\ m \end{matrix} \right|_2 z^{-n-1}$$

is absolutely convergent for all $z \in \mathbb{C}$ such that $|z| > 1$ and, in that region, has sum

$$F_m(z) = \left(\frac{z}{1-z} \frac{d}{dz} \right)^{m-1} \frac{1}{z(z-1)}.$$

The double series

$$\sum_{m=1}^{\infty} \sum_{n=1}^{\infty} \left| \begin{matrix} n \\ m \end{matrix} \right|_2 z^{-n-1} \frac{w^m}{m!}$$

is absolutely convergent in $\{(z, w) \in \mathbb{C}^2 : |z|^{1/2} > 1 + |w|^{1/2}\}$.

Proof. Since $\left| \begin{matrix} n \\ 1 \end{matrix} \right|_2 = 1$ for all $n \in \mathbb{N}$, the case $m = 1$ is immediate. If $m \in \mathbb{N}$ then (working as in the proof of Lemma 11)

$$(9) \quad \binom{n}{m} \frac{(m+1)!}{1!} \leq \left| \begin{matrix} n+1 \\ m+1 \end{matrix} \right|_2 \leq \binom{n}{m} \frac{(n+1)!}{(n+1-m)!} \quad \forall n \geq m$$

and the ratio and comparison tests yield convergence as claimed. Furthermore,

$$\begin{aligned} F_{m+1}(z) &:= \sum_{n=m+1}^{\infty} \left| \begin{matrix} n \\ m+1 \end{matrix} \right|_2 z^{-n-1} = \sum_{n=m+1}^{\infty} \left(\left| \begin{matrix} n-1 \\ m+1 \end{matrix} \right|_2 + n \left| \begin{matrix} n-1 \\ m \end{matrix} \right|_2 \right) z^{-n-1} \\ &= \sum_{l=m}^{\infty} \left| \begin{matrix} l \\ m+1 \end{matrix} \right|_2 z^{-l-2} - \sum_{n=m+1}^{\infty} \left| \begin{matrix} n-1 \\ m \end{matrix} \right|_2 \frac{d}{dz} z^{-n} \\ &= z^{-1} F_{m+1}(z) - F'_m(z) \end{aligned}$$

and so

$$F_{m+1}(z) = \frac{z}{1-z} F'_m(z) = \dots = \left(\frac{z}{1-z} \frac{d}{dz} \right)^m F_1(z) \quad (|z| > 1).$$

For the final claim, note that, by (9), if $m \in \mathbb{N}$ and $|z| > 1$ then

$$\begin{aligned} \sum_{n=1}^N \binom{n}{m} |z|^{-n-1} \frac{|w|^m}{m!} &\leq \sum_{n=m}^N \binom{n}{m}^2 \frac{m}{n(n+1-m)} |z|^{-n-1} |w|^m \\ &\leq \left(\sum_{n=m}^{\infty} \binom{n}{m} |z|^{-(n+1)/2} \right)^2 |w|^m = (|z|^{1/2} - 1)^{-2(m+1)} |w|^m \end{aligned}$$

for all $m \geq N$; the last step follows from (6). The result follows by the comparison test. ■

DEFINITION 15. The *Lambert W function* (also known as the *product logarithm*) is the multi-valued inverse to the map $w \mapsto we^w$. Following the conventions of Corless *et al.* [8], the branch

$$\widetilde{W}: z \mapsto \begin{cases} W_{-1}(z) & \text{if } z = -e^{-1} + re^{i\theta} \quad (r > 0, \theta \in]0, \pi]) \\ W_0(z) & \text{if } z = -e^{-1} + re^{i\theta} \quad (r \geq 0, \theta \in]\pi, 2\pi]) \end{cases}$$

is holomorphic on $\mathbb{C} \setminus [-e^{-1}, \infty[$ and has range containing $\Pi_- := \{z \in \mathbb{C} : \text{Im } z \in]-\pi, 0]\}$.

THEOREM 16. *If X is a random variable with monotone Poisson distribution and mean $t > 0$, the Cauchy transform of X equals*

$$z \mapsto \frac{-1}{\widetilde{W}(-ze^{t-z})} \quad \forall z \in \Pi_+ := \{z \in \mathbb{C} : \text{Im } z \in]0, \pi]\},$$

where \widetilde{W} is the branch of the Lambert W function given in Definition 15.

Proof. The function

$$\begin{aligned} g: \text{Dom } g := \{x + iy : x \in \mathbb{R}, |y| < \pi\} &\rightarrow \{x \cot x + iy : |y| < x < \pi\} \setminus [0, 1]; \\ z &\mapsto -W_0(e^z) \end{aligned}$$

is a holomorphic bijection and

$$g'(z) = -e^z W'_0(e^z) = \frac{-e^z e^{-W_0(e^z)}}{1 + W_0(e^z)} = \frac{-W_0(e^z)}{1 - (-W_0(e^z))} = \frac{g(z)}{1 - g(z)} \quad \forall z \in \text{Dom } g.$$

If $h(z) := z/(1-z)$ then $h \circ g = g'$ and the chain rule implies that

$$\frac{d}{dz} \circ c_g = m_{g'} \circ c_g \circ \frac{d}{dz} \Leftrightarrow \frac{d}{dz} \circ c_g = c_g \circ m_h \circ \frac{d}{dz} \Leftrightarrow c_{g^{-1}} \circ \frac{d}{dz} \circ c_g = m_h \circ \frac{d}{dz},$$

where $m_h: f \mapsto fh$ and $c_g: f \mapsto f \circ g$; the above holds when applied to any holomorphic function with domain including the range of g (which does not include 0 or 1). Hence, by Lemma 14,

$$F'_m(z) = \left(\frac{z}{1-z} \frac{d}{dz} \right)^{m-1} \frac{1}{z(z-1)} = \frac{d^{m-1}}{dw^{m-1}} \frac{1}{g(w)(g(w)-1)} \Big|_{w=g^{-1}(z)} = (1/g)^{(m)}(g^{-1}(z))$$

for all $z \in \text{Ran } g$ such that $|z| > 1$; note that

$$\frac{1}{g(z)(g(z)-1)} = -g'(z)/g(z)^2 = (1/g)'(z) \quad \forall z \in \text{Dom } g.$$

Fix $z \in \text{Ran } g$ such that $|z|^{1/2} > 1 + t^{1/2}$ and let $\varepsilon > 0$ be such that $g^{-1}(z) + w \in \text{Dom } g$ for all $w \in \mathbb{C}$ such that $|w| < \varepsilon$; suppose further that $\varepsilon^{1/2} < |z|^{1/2} - 1$. By Taylor's theorem and Lemma 14, if $|w| < \varepsilon$ then

$$\frac{1}{g(g^{-1}(z) + w)} = \frac{1}{z} + \sum_{m=1}^{\infty} (1/g)^{(m)}(g^{-1}(z)) \frac{w^m}{m!} = \frac{1}{z} + \sum_{m=1}^{\infty} \sum_{n=1}^{\infty} \left| \begin{matrix} n \\ m \end{matrix} \right|_2 z^{-n-1} \frac{w^m}{m!}.$$

As the first and last functions above are holomorphic functions of w in

$$\{w \in \mathbb{C} : g^{-1}(z) + w \in \text{Dom } g\} \quad \text{and} \quad \{w \in \mathbb{C} : |w|^{1/2} < |z|^{1/2} - 1\},$$

respectively, and agree on $\{w \in \mathbb{C} : |w| < \varepsilon\}$, which lies in their intersection, they agree at t . Thus

$$\frac{-1}{W_0(-ze^{t-z})} = \frac{1}{g(g^{-1}(z) + t)} = \frac{1}{z} + \sum_{m=1}^{\infty} \sum_{n=1}^{\infty} \left| \begin{matrix} n \\ m \end{matrix} \right|_2 z^{-n-1} \frac{t^m}{m!} = G_X(z)$$

for all $z \in \text{Ran } g$ such that $|z|^{1/2} > 1 + t^{1/2}$, where the exchange of the order of summation which yields the last equality is a consequence of absolute convergence. Finally, G_X is holomorphic on $\mathbb{C} \setminus \mathbb{R}$ and $z \mapsto -1/\widetilde{W}(-ze^{t-z})$ is holomorphic on Π_+ ; a short calculation shows that if $z \in \Pi_+$ then $-ze^{t-z} \in \text{Dom } \widetilde{W}$ and

$$\frac{-1}{\widetilde{W}(-ze^{t-z})} = \frac{-1}{W_0(-ze^{t-z})} \quad \forall z \in \text{Ran } g \cap \{z \in \mathbb{C} : \text{Im } z > 0\}.$$

Since G_X and $z \mapsto -1/W_0(-ze^{t-z})$ agree on $\{z \in \text{Ran } g : |z|^{1/2} > 1 + t^{1/2}\}$ and this set meets Π_+ , the result follows. ■

The following is contained in [16, Theorem 4.1] and may be proved, as it is there, by employing the Stieltjes inversion formula; the fact that the monotone Poisson distribution has an atom only at the origin will be useful later.

COROLLARY 17. *If X is a random variable distributed according to the monotone Poisson law with mean $t > 0$ then*

$$\mathbb{P}(X \in A) = e^{-t} \delta_0(A) + \frac{1}{\pi} \int_{A \cap]a, b[} \text{Im} \frac{1}{W_{-1}(-xe^{t-x})} dx$$

for any measurable set $A \subseteq \mathbb{R}$, where δ_0 is the Dirac measure which is supported at 0, $a = -W_0(-e^{-(1+t)})$ and $b = -W_{-1}(-e^{-(1+t)})$.

5. The classical process. It is well known [3, Section 4.4] that monotone Brownian motion corresponds, *via* the isomorphism between quantum semimartingale algebras, to Azéma's martingale, i.e., the classical normal martingale $m = (m_t)_{t \geq 0}$ which satisfies the structure equation

$$m_0 = 0, \quad d[m]_t = -m_{t-} dm_t + dt.$$

A similar correspondence holds for the monotone Poisson process.

THEOREM 18. *The quantum stochastic differential equation*

$$(10) \quad Y_0 = 0, \quad dY = (I - Y) d\Lambda + dA + dA^\dagger + dt$$

has a unique solution in the class of HP-adapted processes acting on exponential domains; this solution consists of commuting, self-adjoint operators such that $\|Y_t\| \leq (1 + \sqrt{t})^2$ and, in the vacuum state, Y_t obeys the monotone Poisson law with mean t , for all $t \geq 0$.

Proof. If $P = (P_t)_{t \geq 0}$ is the Ω -adapted process such that

$$P_t = \int_0^t \mathbb{E}_s(d\Lambda_s + dA_s + dA_s^\dagger + ds) \quad \forall t \geq 0,$$

where \mathbb{E} is the conditional-expectation process in Boson Fock space, then $Y := \hat{\pi}(P)$, the HP-adapted projection of P , satisfies (10) [3, Corollary 31]; alternatively, Y can be viewed as the image of the Poisson process under Attal's \mathcal{D} transform [2]. As P is clearly bounded and self-adjoint, so is Y , and

$$\|Y_t\| = \|P_t\| \leq \|\mathbb{E}\|_{\infty,t} + 2\|\mathbb{E}\|_{2,t} + \|\mathbb{E}\|_{1,t} = 1 + 2\sqrt{t} + t \quad \forall t \geq 0.$$

Since Y_t and P_t have the same moments in the vacuum state, [4, Corollary 4.1] gives the claim about the distribution of Y_t .

For uniqueness, if X is a solution of (10) which acts on some exponential domain \mathcal{E}_0 then

$$(X - Y)_0 = 0, \quad d(X - Y) = -(X - Y) d\Lambda \quad \text{on } \mathcal{E}_0;$$

as is well known, this equation admits only the zero process as a solution. Finally, if $s \geq 0$ is fixed and $X_t := Y_{\min\{s,t\}}$ for all $t \geq 0$ then the quantum Itô formula yields

$$d(XY - YX) = -(XY - YX) d\Lambda,$$

so $(XY - YX)_t = 0$ for all $t \geq 0$ and commutativity follows. ■

LEMMA 19. *If M and N are regular quantum semimartingales and $t \geq 0$ then*

$$\int_0^t M_s dN_s = \lim_{n \rightarrow \infty} \sum_{i=1}^n M_{t_{i-1}^n} (N_{t_i^n} - N_{t_{i-1}^n})$$

strongly on exponential vectors, where $0 = t_0^n < t_1^n < \dots < t_n^n = t$ for all $n \geq 1$ are such that $\sup\{t_i^n - t_{i-1}^n : i = 1, \dots, n\} \rightarrow 0$ as $n \rightarrow \infty$.

Proof. If $dN = E d\Lambda + F dA + G dA^\dagger + H dt$, $M_t^{(n)} := M_{t_{i-1}^n}$ for $t \in [t_{i-1}^n, t_i^n[$ (where $i = 1, \dots, n$) and $f \in L^2(\mathbb{R}_+)$ then [14, VI.1(6.7),(7.4),(7.8)] there exists a constant C_f such that, for all $t \geq 0$,

$$\begin{aligned} & \left\| \int_0^t (M_s - M_s^{(n)}) dN_s \varepsilon(f) \right\|^2 \\ & \leq C_f \int_0^t (|f(s)|^2 \| (M_s - M_s^{(n)}) E_s \varepsilon(f) \|^2 + \| (M_s - M_s^{(n)}) F_s \varepsilon(f) \|^2 \\ & \quad + \| (M_s - M_s^{(n)}) G_s \varepsilon(f) \|^2) ds + 2 \left(\int_0^t \| (M_s - M_s^{(n)}) H_s \varepsilon(f) \|^2 ds \right)^2 \rightarrow 0 \end{aligned}$$

as $n \rightarrow \infty$, by the dominated-convergence theorem; recall that $\mathbb{R}_+ \ni t \mapsto M_t$ is strongly continuous on the whole of Fock space. The result follows because

$$X \int_s^t dN_r \varepsilon(f) = \int_s^t X dN_r \varepsilon(f) \quad \forall X \in \mathcal{B}(\mathfrak{H}_s)$$

if $0 \leq s \leq t$ (where \mathfrak{H}_s denotes Boson Fock space up to time s); this is immediate when E, F, G and H are step functions and approximation with such yields the general case. ■

LEMMA 20. *If $(\Omega, \mathcal{F}, \mathbb{P})$ is a probability space, $(z_t : \Omega \rightarrow \mathbb{R})_{t \geq 0}$ a family of bounded, \mathcal{F} -measurable functions and $(\mathcal{F}_t := \sigma(z_s : 0 \leq s \leq t))_{t \geq 0}$ their natural filtration then $\text{alg}\{1, z_s : 0 \leq s \leq t\}$ is dense in $L^2(\mathcal{F}_t)$ for all $t \geq 0$.*

Proof. If $f \in L^2(\mathcal{F}_t)$ then there exists $(t_n)_{n \geq 1}$ such that $f \in L^2(\sigma(z_{t_n} : n \geq 1))$; let $\mathcal{G}_n := \sigma(z_{t_m} : 1 \leq m \leq n)$ and note that $\mathbb{E}[f|\mathcal{G}_n] \rightarrow f$ as $n \rightarrow \infty$ (almost surely and in the L^2 sense). Hence it suffices to prove that $\mathfrak{A}_n := \text{alg}\{1, z_{t_m} : 1 \leq m \leq n\}$ is dense in $L^2(\mathcal{G}_n)$; fix $(\alpha_1, \dots, \alpha_n) \in \mathbb{R}^n$ and let $(p_k)_{k \geq 1}$ be a sequence of polynomials such that $p_k(x_1, \dots, x_n) \rightarrow \exp(-i(\alpha_1 x_1 + \dots + \alpha_n x_n))$ locally uniformly on \mathbb{R}^n . If $g \in L^2(\mathcal{G}_n)$ is orthogonal to \mathfrak{A}_n then $g = h(z_{t_1}, \dots, z_{t_n})$ for some Borel-measurable function $h : \mathbb{R}^n \rightarrow \mathbb{C}$ and

$$0 = \mathbb{E}[gp_k(z_{t_1}, \dots, z_{t_n})] \rightarrow \mathbb{E}[g \exp(-i(\alpha_1 z_{t_1} + \dots + \alpha_n z_{t_n}))] = \int_{\mathbb{R}^n} h(\mathbf{x}) \exp(-i\boldsymbol{\alpha} \cdot \mathbf{x}) d\nu(\mathbf{x}),$$

where ν is the distribution of $(z_{t_1}, \dots, z_{t_n})$; the result follows from the injectivity of the Fourier transform. ■

THEOREM 21. *There exists a martingale $x = (x_t)_{t \geq 0}$ (with respect to a filtration satisfying the usual conditions) which satisfies the structure equation*

$$(11) \quad x_0 = 0, \quad d[x]_t = (1 - t - x_{t-}) dx_t + dt;$$

this process x is unique in law and has the chaotic-representation property.

Proof. Identify \mathfrak{A} , the unital C^* algebra generated by the family of operators $\{Y_t : t \geq 0\}$ from Theorem 18, with $C(\Delta)$, where Δ is a compact Hausdorff space, via the Gelfand map $A \mapsto \widehat{A}$. If Ω denotes the vacuum vector then $\widehat{A} \mapsto \langle \Omega, A\Omega \rangle$ is a positive linear functional on $C(\Delta)$ and the Riesz-Markov-Kakutani theorem yields a unique regular Borel measure μ on Δ such that

$$\langle \Omega, A\Omega \rangle = \int_{\Delta} \widehat{A} d\mu \quad \forall A \in \mathfrak{A};$$

since $\|\Omega\| = 1$, μ is a probability measure. Thus if $X_t := Y_t - t$ then $(x_t^\circ := \widehat{X}_t)_{t \geq 0}$ is a family of square-integrable random variables on the probability space $(\Delta, \mathcal{B}, \mu)$ (where \mathcal{B} is the Borel σ -algebra of Δ).

Let \mathcal{F}_t° denote the σ -algebra generated by $\{x_s^\circ : 0 \leq s \leq t\}$, so that $(x_t^\circ)_{t \geq 0}$ is adapted to the filtration $(\mathcal{F}_t^\circ)_{t \geq 0}$. Since $\mathbb{E}_s(Y_t - t)\mathbb{E}_s = (Y_s - s)\mathbb{E}_s$ and $\mathbb{E}_s Y_r \mathbb{E}_s = Y_r \mathbb{E}_s$ for all $t \geq s \geq r \geq 0$, if $\widehat{Z} = z \in \text{alg}\{1, x_r^\circ : 0 \leq r \leq s\}$ then $\mathbb{E}_s Z \mathbb{E}_s = Z \mathbb{E}_s$ and

$$\int_{\Delta} \bar{z}(x_t^\circ - x_s^\circ) d\mu = \langle Z\Omega, (Y_t - t - Y_s + s)\Omega \rangle = \langle Z\Omega, \mathbb{E}_s(Y_t - t - Y_s + s)\mathbb{E}_s\Omega \rangle = 0$$

if $t \geq s \geq 0$. Thus (by Lemma 20) $\mathbb{E}[x_t^\circ | \mathcal{F}_s^\circ] = x_s^\circ$ and $(x_t^\circ)_{t \geq 0}$ is a $(\mathcal{F}_t^\circ)_{t \geq 0}$ -martingale such that $\mathbb{E}[(x_t^\circ)^2] = \langle \Omega, X_t^2 \Omega \rangle = t$ for all $t \geq 0$, by the quantum Itô formula (or (1)). Furthermore, $\mathbb{E}[(x_t^\circ - x_s^\circ)^2] = t - s$ so $t \mapsto x_t^\circ$ is L^1 -continuous; in particular, it has a modification $(x_t)_{t \geq 0}$ with càdlàg paths [18, Proposition II.67.6 and Theorem II.66.2] which is a martingale in the filtered probability space $(\Omega, \mathcal{F}, \mathbb{P}, (\mathcal{F}_t)_{t \geq 0})$, the μ -completion

of $(\Delta, \mathcal{B}, \mu)$ equipped with the usual augmentation of $(\mathcal{F}_t^c)_{t \geq 0}$ (which is right continuous and such that \mathcal{F}_0 is \mathbb{P} -complete) [18, Definition II.67.3].

The quantum Itô formula also implies that $X_t^2 = X_t + \int_0^t (X_s - s) dX_s + t$ for all $t \geq 0$, so if $z = \widehat{Z} \in \widehat{\mathfrak{A}} = C(\Delta)$ and $t \geq 0$ then, by Lemma 19,

$$\begin{aligned} \int_{\Delta} \bar{z}(x_t^2 - x_t - t) d\mathbb{P} &= \langle Z\Omega, (X_t^2 - X_t - t)\Omega \rangle \\ &= \lim_{n \rightarrow \infty} \langle Z\Omega, \sum_{i=1}^n (X_{t_{i-1}^n} - t_{i-1}^n)(X_{t_i^n} - X_{t_{i-1}^n})\Omega \rangle \\ &= \lim_{n \rightarrow \infty} \int_{\Delta} \bar{z} \sum_{i=1}^n (x_{t_{i-1}^n} - t_{i-1}^n)(x_{t_i^n} - x_{t_{i-1}^n}) d\mathbb{P} \\ &= \int_{\Delta} \bar{z} \int_0^t (x_{s-} - s) dx_s d\mathbb{P}; \end{aligned}$$

since $C(\Delta)$ is dense in $L^2(\mathbb{P})$, $x_t^2 - x_t - t = \int_0^t (x_{s-} - s) dx_s$ and $(x_t)_{t \geq 0}$ is a normal martingale which satisfies the structure equation as claimed.

The other results follow from [9, Lemma 7] (with $f(t) = 1 - t$ and $g(s, t) = -1$ for all $s, t \geq 0$), working as in [9, Proof of Proposition 6]. ■

REMARK 22. The existence of a solution to (11) also follows from a much more general result of Taviot [20, Chapitre 4], which uses completely classical methods.

LEMMA 23. *The martingale x is purely discontinuous.*

Proof. Note that, as $x_t^c = \int_0^t 1_{\{x_{s-} = 1-s\}} dx_s$ [9, Proposition 1(ii)],

$$\mathbb{E}[x_t^c] = \mathbb{E}\left[\int_0^t 1_{\{x_{s-} = 1-s\}} ds\right] = \int_0^t \mathbb{P}(x_s = 1 - s) ds = 0;$$

this holds since $\{s : x_{s-}(\omega) \neq x_s(\omega)\}$ is countable almost surely and $\mathbb{P}(x_s = y)$ is non-zero if and only if $y = -s$ (note that $x_s + s$ is distributed according to the monotone Poisson law, which is atomic only at 0 by Corollary 17). Thus $x^c \equiv 0$, as claimed. ■

THEOREM 24. *If $(x_t)_{t \geq 0}$ is as in Theorem 21 and $y_t := x_t + t$ then $(y_t)_{t \geq 0}$ is such that $\Delta y_t := y_t - y_{t-} \in \{0, 1 - y_{t-}\}$ for all $t > 0$. If*

$$g_t := \sup\{s \in [0, t] : y_s = 1\} \in \{-\infty\} \cup]0, t[$$

is the last time up to t that y hits 1 then

$$(12) \quad y_t = -W_{\bullet}(-\exp(-1 - t + g_t)) \quad \forall t \geq 0,$$

where W_{\bullet} equals W_0 or W_{-1} (Definition 15), so $y_t \in [0, 1]$ or $y_t \in [1, \infty[$, respectively.

Proof. The statement about jumps follows from [9, Proposition 1(i)], or directly from the equation

$$d[y]_t = (1 - y_{t-}) dy_t + y_{t-} dt;$$

this (and the identity $y_t^2 = 2 \int_0^t y_{s-} dy_s + [y]_t$) also implies that

$$(y_t - 1)^2 = 1 + 2 \int_0^t y_{s-} ds - [y]_t \quad \forall t \geq 0.$$

As (12) is clear for $t = 0$, let $t > 0$ and fix $\omega \in \Omega$ (which will be omitted in what follows, for the sake of brevity). If $g_t = -\infty$ then $y_s \neq 1$ for all $s \in [0, t]$, so $\Delta y_s = 0$ for all $s \in]0, t]$ and y is continuous on $[0, t]$. As x is purely discontinuous, by Lemma 23,

$$[y]_t = [x]_t = \sum_{0 < s \leq t} \Delta x_s^2 = \sum_{0 < s \leq t} \Delta y_s^2 = 0,$$

whence

$$(y_s - 1)^2 = 1 + 2 \int_0^s y_r \, dr \quad \forall s \in [0, t].$$

If $s \in [0, t]$ then $y_s \neq 1$, so $y_r + y_s \neq 2$ for all r sufficiently near s and

$$\frac{y_r - y_s}{r - s} = \frac{(y_r - 1)^2 - (y_s - 1)^2}{(y_r + y_s - 2)(r - s)} = \frac{2}{(y_r + y_s - 2)(r - s)} \int_s^r y_u \, du \rightarrow \frac{y_s}{y_s - 1}$$

as $r \rightarrow s$. Hence $dy_s/ds = y_s/(y_s - 1)$ for all $s \in [0, t]$ (taking one-sided derivatives at $s = 0$ and $s = t$); from the initial condition $y_0 = 0$ it follows that $y \equiv 0$ on $[0, t]$, by Lemma 25, as required.

If $g_t \in]0, t]$ then it is readily verified that $y_{g_t} = 1$; if $g_t = t$ then (12) holds, so suppose that $g_t < t$. This implies that $y_s \neq 1$ for all $s \in]g_t, t]$ and y is continuous on $]g_t, t]$; similar working to the above gives that $dy_s/ds = y_s/(y_s - 1)$ for all $s \in]g_t, t]$, and the result follows from Lemma 25 (working on $[g_t + 1/n, t]$ and letting $n \rightarrow \infty$; note that $y_{g_t + 1/n} > 0$, since $y_{g_t} = 1$, for sufficiently large n). ■

LEMMA 25. *If $f : [s, t] \rightarrow \mathbb{R} \setminus \{1\}$ is continuously differentiable, with $f(s) = b \geq 0$ and $f'(r) = f(r)/(f(r) - 1)$ for all $r \in [s, t]$ (taking one-sided derivatives at $r = s$ and $r = t$), then*

$$(13) \quad f(r) = -W_\bullet(-b \exp(-r + s - b)) \quad \forall r \in [s, t],$$

where $\bullet = 0$ if $b \in [0, 1[$ and $\bullet = -1$ if $b \in]1, \infty[$.

Proof. Let $g(r)$ denote the right-hand side of (13); it is readily verified that $g(s) = b$, $g(r) \neq 1$ and $g'(r) = g(r)/(g(r) - 1)$ for all $r \in [s, t]$. If $c := \sup\{|f(r) - 1|^{-1} : s \leq r \leq t\}$ and $d := \sup\{|g(r) - 1|^{-1} : s \leq r \leq t\}$ then the fundamental theorem of calculus implies that

$$\begin{aligned} |(f - g)(r)| &\leq cd \int_s^r |(f - g)(u)| \, du \leq \dots \\ &\leq \frac{(cd)^n (r - s)^n}{n!} \sup\{|(f - g)(u)| : s \leq u \leq r\} \rightarrow 0 \end{aligned}$$

as $n \rightarrow \infty$, for all $r \in [s, t]$, which gives the result. ■

THEOREM 26. *For all $h > 0$ there exists a family of random variables $(a_{h,n})_{n \in \mathbb{Z}_+}$ such that $a_{h,0} = 0$,*

$$a_{h,n+1}^2 = h + \left(1 - nh - \sum_{k=0}^n a_{h,k}\right) a_{h,n+1} \quad \text{and} \quad \mathbb{E}[a_{h,n+1} | \sigma(a_{h,k} : 0 \leq k \leq n)] = 0.$$

If $w_{h,t} := \sum_{n=0}^{\lfloor t/h \rfloor} a_{h,n}$ for all $t \geq 0$, where $\lfloor s \rfloor$ denotes the greatest integer less than or equal to s , then $w_h := (w_{h,t})_{t \geq 0}$ is a martingale with respect to its natural filtration and

satisfies

$$w_{h,0} = 0, \quad d[w_h]_t = (1 - t - w_{h,t-}) dw_{h,t} + h d[t/h].$$

There exists a sequence $(h_n)_{n \geq 0}$ such that $(w_{h_n,t})_{t \geq 0}$ converges weakly to the solution of the structure equation (11).

Proof. If such a family $(a_{h,n})_{n \in \mathbb{Z}_+}$ were to exist and $\phi_n := (1 - nh - w_{h,nh})/2$ then $w_{h,(n+1)h} = w_{h,nh} + \phi_{h,n} + \varepsilon_{h,n} \sqrt{\phi_{h,n}^2 + h}$ for all $n \geq 0$, where

$$\mathbb{P}(\varepsilon_{h,n} = \pm 1 | w_{h,nh}) = \frac{1}{2} \left(1 \mp \frac{\phi_{h,n}}{\sqrt{\phi_{h,n}^2 + h}} \right),$$

so $(w_{h,nh})_{n \in \mathbb{Z}_+}$ can be constructed as a non-homogeneous Markov process with these transition probabilities. The claims about w_h are simple to verify; it is straightforward to check that

$$w_{h,t}^2 = w_{h,t} + \int_0^t (w_{h,s-} - s) dw_{h,s} + [t/h]h \quad \forall t \geq 0.$$

Finally, that $(w_{h_n,t})_{t \geq 0}$ converges to a solution of (11) for some $(h_n)_{n \geq 0}$ may be proved using the same techniques as Meyer [13]. ■

REMARK 27. Using results of Attal [1] on representable probabilistic interpretations of Fock space, it can be shown that the classical solution constructed in Theorem 26 gives rise to the process that satisfies the quantum stochastic differential equation (10).

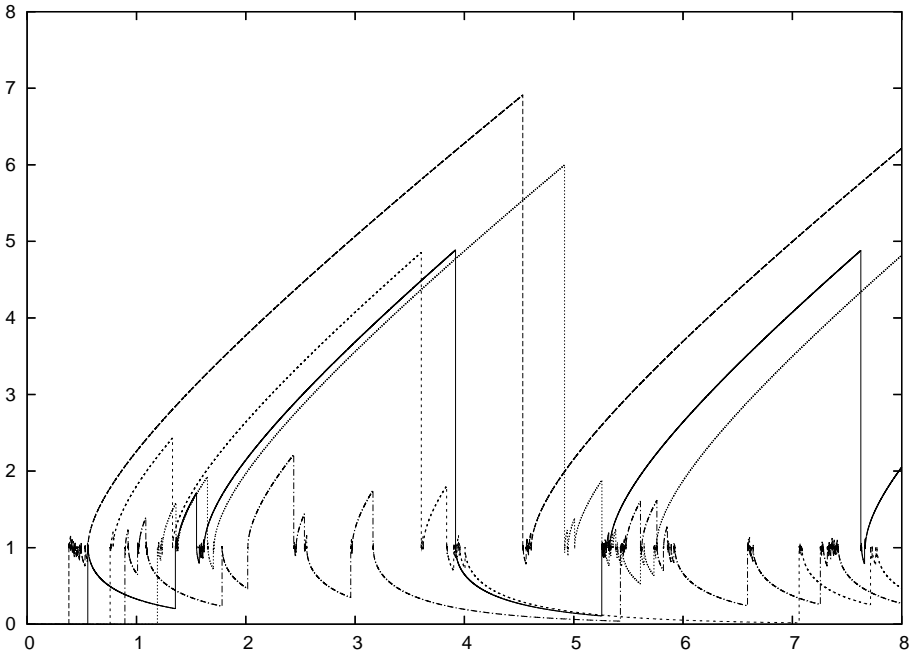


Fig. 1. Approximate sample paths of the monotone Poisson process

REMARK 28. The calculation in the proof of Theorem 26 allows a numerical simulation to be performed: Figure 1 shows some sample paths of

$$t \mapsto \tilde{y}_{h,t} := \sum_{n=0}^{\lfloor t/h \rfloor} a_{h,n} + (t/h - \lfloor t/h \rfloor)a_{h,\lfloor t/h \rfloor+1} + t,$$

with $h = 0.005$ and $t \in [0, 8[$.

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