

## Approximation of a symmetric $\alpha$ -stable Lévy process by a Lévy process with finite moments of all orders

by

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**Abstract.** In this paper we consider a symmetric  $\alpha$ -stable Lévy process  $Z$ . We use a series representation of  $Z$  to condition it on the largest jump. Under this condition,  $Z$  can be presented as a sum of two independent processes. One of them is a Lévy process  $Y_x$  parametrized by  $x > 0$  which has finite moments of all orders. We show that  $Y_x$  converges to  $Z$  uniformly on compact sets with probability one as  $x \downarrow 0$ . The first term in the cumulant expansion of  $Y_x$  corresponds to a Brownian motion which implies that  $Y_x$  can be approximated by Brownian motion when  $x$  is large. We also study integrals of a non-random function with respect to  $Y_x$  and derive the covariance function of those integrals. A symmetric  $\alpha$ -stable random vector is approximated with probability one by a random vector with components having finite second moments.

**1. Introduction.** Stable Lévy processes play an important role among stable processes, similar to that of Brownian motion among Gaussian processes. Thus we start with the definition of a symmetric  $\alpha$ -stable Lévy process (for a comprehensive treatment of stable variables and processes see e.g. Janicki and Weron [2] or Samorodnitsky and Taqqu [8]). We will consider stochastic processes on the time interval  $[0, 1]$ .

**DEFINITION 1.1.** A stochastic process  $\{Z(t), 0 \leq t \leq 1\}$  is called a *symmetric  $\alpha$ -stable Lévy process* ( $0 < \alpha \leq 2$ ) if

1.  $Z(0) = 0$  a.s.
2.  $Z$  has independent increments.
3.  $Z(t) - Z(s)$  has distribution  $S_\alpha(\sigma(t-s)^{1/\alpha}, 0, 0)$  for any  $0 \leq s < t \leq 1$ , that is,  $\alpha$ -stable distribution with scale parameter  $\sigma(t-s)^{1/\alpha}$ , and skewness and shift parameters equal to zero.

For  $\alpha = 2$  we get Brownian motion. If the first two assumptions are satisfied and the process has stationary increments we call it a *Lévy process*. In this paper we will consider  $0 < \alpha < 2$ .

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The *Lévy–Itô integral representation* for a symmetric Lévy process  $Y$  is the following:

$$(1) \quad Y(t) = \int_{\mathbb{R} \setminus 0} y N_t(dy)$$

where  $N$  is the point process of jumps of  $Y$ :  $N = \sum_{s: \Delta Y(s) \neq 0} \delta_{(s, \Delta Y(s))}$  (see e.g. Kallenberg [3]).  $N$  is a Poisson point process with the mean measure  $ds \times \nu(dy)$  where  $\nu(dy)$  is a Lévy measure on  $\mathbb{R} \setminus 0$ . Series representations for Lévy processes can be derived from the Lévy–Itô integral representation (see Rosiński [7]). The series representation for  $\alpha$ -stable Lévy processes is called LePage’s representation (see LePage [4] or Samorodnitsky and Taqqu [8]).

**2. The series representation of a symmetric  $\alpha$ -stable Lévy process.** A symmetric  $\alpha$ -stable Lévy process can be represented as the following series:

$$(2) \quad Z(t) = \sigma C_\alpha^{1/\alpha} \sum_{k=1}^{\infty} \Gamma_k^{-1/\alpha} \gamma_k \mathbf{I}\{U_k \leq t\}$$

where  $0 \leq t \leq 1$  and  $\{\Gamma_k\}_{k=1}^{\infty}$  is a sequence of arrival epochs in a Poisson process with unit arrival rate,  $\{\gamma_k\}_{k=1}^{\infty}$  is a sequence of iid random variables satisfying

$$\mathbb{P}(\gamma_k = 1) = \mathbb{P}(\gamma_k = -1) = 1/2,$$

and  $\{U_k\}_{k=1}^{\infty}$  is a sequence of iid random variables uniformly distributed on  $[0, 1]$ . These sequences are independent and

$$(3) \quad C_\alpha = \left( \int_0^\infty s^{-\alpha} \sin s \, ds \right)^{-1} = \begin{cases} \frac{1-\alpha}{\Gamma(2-\alpha) \cos(\pi\alpha/2)} & \text{if } \alpha \neq 1, \\ 2/\pi & \text{if } \alpha = 1. \end{cases}$$

A symmetric  $\alpha$ -stable Lévy process can be regarded as a pure jump process. The epochs of its jumps are described by  $\{U_k\}_{k=1}^{\infty}$ , the direction of a jump is governed by  $\{\gamma_k\}_{k=1}^{\infty}$  and the heights of the jumps are presented in decreasing order by  $\{\sigma C_\alpha^{1/\alpha} \Gamma_k^{-1/\alpha}\}_{k=1}^{\infty}$ .

Let us introduce the Poisson random measure and the Lévy measure generated by the process  $Z$ . Define

$$(4) \quad N_t(A) = N([0, t], A) = \sum_{0 \leq s \leq t} \mathbf{I}_A(\Delta Z_s)$$

for  $0 \leq t \leq 1$  and any Borel subset  $A$  of  $\mathbb{R}$  such that  $0 \notin \bar{A}$ . The process  $N_t(A)$ ,  $0 \leq t \leq 1$ , is a Poisson process with arrival rate

$$\nu(A) = \mathbb{E}N_1(A).$$

The measure  $\nu$  is called the *Lévy measure* of the Lévy process  $Z$ . In fact the measure  $N([0, t], A)$  can be extended to a Poisson measure on  $[0, 1] \times \mathbb{R} \setminus 0$

with the mean measure  $ds \times \nu(dy)$ . Thus using the series representation we obtain

$$(5) \quad N_t(A) = \sum_{k=1}^{\infty} \mathbf{I}_A \{ \sigma C_\alpha^{1/\alpha} \Gamma_k^{-1/\alpha} \gamma_k \} \mathbf{I} \{ U_k \leq t \}$$

and

$$(6) \quad \nu(A) = \sum_{k=1}^{\infty} \mathbb{E} \mathbf{I}_A \{ \sigma C_\alpha^{1/\alpha} \Gamma_k^{-1/\alpha} \gamma_k \}.$$

This Lévy measure has the following form:

$$(7) \quad \nu(dy) = \frac{P}{y^{1+\alpha}} \mathbf{I}_{(0,\infty)}(y) dy + \frac{P}{|y|^{1+\alpha}} \mathbf{I}_{(-\infty,0)}(y) dy$$

where

$$P = \frac{1}{2} \alpha \sigma^\alpha C_\alpha.$$

**3. Conditioning on the largest jump.** Consider a symmetric  $\alpha$ -stable Lévy process under the condition  $\Gamma_1 = x$  where  $x > 0$  ( $\Gamma_1$  has exponential distribution with parameter 1). Then we get

$$(8) \quad Z_x(t) \stackrel{d}{=} \sigma C_\alpha^{1/\alpha} x^{-1/\alpha} \gamma_1 \mathbf{I} \{ U_1 \leq t \} \\ + \sigma C_\alpha^{1/\alpha} \sum_{k=1}^{\infty} (\Gamma_k + x)^{-1/\alpha} \gamma_{k+1} \mathbf{I} \{ U_{k+1} \leq t \}.$$

Therefore let us investigate the following process:

$$(9) \quad Y_x(t) = \sigma C_\alpha^{1/\alpha} \sum_{k=1}^{\infty} (\Gamma_k + x)^{-1/\alpha} \gamma_{k+1} \mathbf{I} \{ U_{k+1} \leq t \} \\ \stackrel{d}{=} \sigma C_\alpha^{1/\alpha} \sum_{k=1}^{\infty} (\Gamma_k + x)^{-1/\alpha} \gamma_k \mathbf{I} \{ U_k \leq t \}$$

and note that

$$(10) \quad Z_x(t) \stackrel{d}{=} A_x(t) + Y_x(t)$$

where  $A_x(t) = \sigma C_\alpha^{1/\alpha} x^{-1/\alpha} \gamma_1 \mathbf{I} \{ U_1 \leq t \}$  and the processes  $A_x$  and  $Y_x$  are independent.

**PROPOSITION 3.1.** *The process  $Y_x$  is well-defined, that is, the sum in (9) converges almost surely.*

*Proof.* This is a special case of Proposition 3.3 below. ■

Let us define the following auxiliary process:

$$(11) \quad Y_{x,z}(t) = \sigma C_\alpha^{1/\alpha} \sum_{k=1}^{\infty} (\Gamma_k + x)^{-1/\alpha} \gamma_k \mathbf{I} \{ U_k \leq t \} \mathbf{I} \{ \Gamma_k < z \}.$$

Notice that the series in (11) consists of finite random number of nonzero terms a.s. (by the strong law of large numbers for the sums  $\Gamma_k$ ) and  $Y_{x,z}(t) \rightarrow Y_x(t)$  a.s. as  $z \rightarrow \infty$  for all  $0 \leq t \leq 1$ . The process (11) can be written using the Poisson measure  $N$  defined in (4) in the following way:

$$(12) \quad Y_{x,z}(t) = \int_{A_z} f_x(y) N_t(dy)$$

where  $A_z = (-\infty, -z^{-1/\alpha}) \cup (z^{-1/\alpha}, \infty)$  and (for simplicity we assume here that  $\sigma C_\alpha^{1/\alpha} = 1$ )

$$(13) \quad f_x(y) = \frac{1}{(|y|^{-\alpha} + x)^{1/\alpha}} \operatorname{sgn}(y).$$

It follows that  $Y_{x,z}$  is a Lévy process (see e.g. Protter [6, Th. 36]).

**THEOREM 3.1.** *The process  $Y_x$  defined in (9) is a Lévy process with finite moments of all orders and*

$$(14) \quad Y_x(t) = \int_{\mathbb{R} \setminus 0} f_x(y) N_t(dy)$$

in the sense of the limit of  $Y_{x,z}(t)$  given in (12) as  $z \rightarrow \infty$  a.s.

*Proof.* The stationarity of the increments follows from the fact that

$$\mathbf{I}\{U_k \leq t\} - \mathbf{I}\{U_k \leq s\} = \mathbf{I}\{s < U_k \leq t\} \stackrel{d}{=} \mathbf{I}\{U_k \leq t - s\}$$

for all  $0 \leq s < t \leq 1$ . From the convergence  $Y_{x,z}(t) \rightarrow Y_x(t)$  a.s. as  $z \rightarrow \infty$  for all  $0 \leq t \leq 1$  we obtain the independence of the increments of  $Y_x$ . Since  $Y_x$  has bounded jumps, it has finite moments of all orders (see Protter [6, Th. 34]). ■

We derive the exact form of the Lévy measure of the process  $Y_x$  and the characteristic function of the random variable  $Y_x(t)$ .

**THEOREM 3.2.** *The Lévy measure of the process  $Y_x$  has the form*

$$(15) \quad \nu_x(dy) = \frac{P}{y^{1+\alpha}} \mathbf{I}_{(0, \sigma C_\alpha^{1/\alpha} x^{-1/\alpha})}(y) dy + \frac{P}{|y|^{1+\alpha}} \mathbf{I}_{(-\sigma C_\alpha^{1/\alpha} x^{-1/\alpha}, 0)}(y) dy.$$

*The characteristic function of the random variable  $Y_x(t)$  is*

$$(16) \quad \mathbb{E} \exp(iuY_x(t)) = \exp\left(-t|u|^\alpha 2P \int_0^{|\sigma C_\alpha^{1/\alpha} x^{-1/\alpha}|} \frac{1 - \cos s}{s^{1+\alpha}} ds\right)$$

$$(17) \quad = \exp\left(-t\alpha \sum_{k=1}^{\infty} \frac{(-1)^{k-1}}{(2k - \alpha)(2k)!} (u\sigma C_\alpha^{1/\alpha} x^{-1/\alpha})^{2k}\right).$$

*Proof.* Let  $\nu_x$  be the Lévy measure of the process  $Y_x$ . Then

$$(18) \quad \nu_x(A) = \sum_{k=1}^{\infty} \mathbb{E} \mathbf{I}_A(\sigma C_\alpha^{1/\alpha}(\Gamma_k + x)^{-1/\alpha} \gamma_k)$$

for any  $A \subset \mathbb{R}$  with  $0 \notin \bar{A}$ . For simplicity we assume that  $\sigma C_\alpha^{1/\alpha} = 1$ . Take  $A = (y, \infty)$  where  $y > 0$  and  $y < x^{-1/\alpha}$  because  $(\Gamma_k + x)^{-1/\alpha} < x^{-1/\alpha}$  a.s. Since  $(\Gamma_k + x)^{-1/\alpha} > y$  is equivalent to  $\Gamma_k^{-1/\alpha} > (y^{-\alpha} - x)^{-1/\alpha}$ , using (6), (7) and (18) we get

$$\nu_x(A) = \int_{(y^{-\alpha} - x)^{-1/\alpha}}^{\infty} \frac{P}{z^{1+\alpha}} dz = \int_y^{x^{-1/\alpha}} \frac{P}{w^{1+\alpha}} dw.$$

For  $y < 0$  we proceed similarly. Thus we obtain

$$\nu_x(dy) = \frac{P}{y^{1+\alpha}} \mathbf{I}_{(0, x^{-1/\alpha})}(y) dy + \frac{P}{|y|^{1+\alpha}} \mathbf{I}_{(-x^{-1/\alpha}, 0)}(y) dy.$$

From the Lévy–Khinchin formula we can write

$$\mathbb{E} \exp(iuY_x(t)) = \exp(-t\psi(u))$$

where for  $u \geq 0$ ,

$$\begin{aligned} \psi(u) &= \int_{\mathbb{R} \setminus 0} (1 - e^{iuy}) \nu_x(dy) = \int_{\mathbb{R} \setminus 0} (1 - \cos(uy)) \nu_x(dy) \\ &= \int_0^{x^{-1/\alpha}} (1 - \cos(uy)) \frac{2P}{y^{1+\alpha}} dy = 2Pu^\alpha \int_0^{ux^{-1/\alpha}} \frac{1 - \cos(s)}{s^{1+\alpha}} ds \end{aligned}$$

where in the last equality we have substituted  $s = uy$ . Arguing similarly for  $u < 0$ , we arrive at (16). To obtain the second form of the characteristic function we expand  $\cos s$  in Taylor series and integrate each summand. ■

In the next proposition we derive the covariance structure of the process  $Y_x$ .

**PROPOSITION 3.2.** *Let  $Y_x$  be defined in (9). Then*

$$(19) \quad \mathbb{E}Y_x(t) = 0,$$

$$(20) \quad \mathbb{E}Y_x^2(t) = \mathbb{E} \left( \int_{\mathbb{R} \setminus 0} f_x(y) N_t(dy) \right)^2 = t \int_{\mathbb{R} \setminus 0} f_x^2(y) \nu(dy) = t\sigma^2(x)$$

and

$$(21) \quad \mathbb{E}Y_x(s)Y_x(t) = \sigma^2(x) \min\{s, t\}$$

where

$$\sigma^2(x) = \frac{\alpha \sigma^2 C_\alpha^{2/\alpha}}{(2 - \alpha)x^{(2-\alpha)/\alpha}}.$$

*Proof.* The expectation is zero because the process  $Y_x$  is symmetric. The form of the variance follows from the well known fact that for Lévy processes without Gaussian component,  $\mathbf{Var} Y_x(t) = t \int y^2 \nu_x(dy)$ . (Twice differentiating the characteristic function in the form (17) also gives the second moment of the random variable  $Y_x(t)$  in the form (20).) To get the desired form of covariance function we use the equality  $Y(s)Y(t) = Y(s)(Y(t) - Y(s)) + Y^2(s)$  for  $s < t$ . ■

REMARK 3.1. Taking only the first term in the series (17) we obtain the characteristic function of Brownian motion  $B_x$  with variance  $\sigma^2(x) = \frac{\alpha\sigma^2 C_\alpha^{2/\alpha}}{(2-\alpha)x^{(2-\alpha)/\alpha}}$  the same as that of  $Y_x$ , that is,

$$\mathbb{E} \exp(iuB_x(t)) = \exp\left(-\frac{tu^2}{2} \frac{\alpha\sigma^2 C_\alpha^{2/\alpha}}{(2-\alpha)x^{(2-\alpha)/\alpha}}\right).$$

Now we investigate the integral of a non-random measurable function  $g$  with respect to the process  $Y_x$ , that is,

$$(22) \quad \int_0^1 g(s) dY_x(s).$$

We impose assumptions on  $g$  to ensure that the second moment of the integral is finite. The integral of a non-random function  $g$  with respect to a symmetric  $\alpha$ -stable Lévy process and the process  $Y_x$  is well defined if  $g \in L^\alpha([0, 1], ds)$  (this follows from Urbanik and Woyczyński [9]).

THEOREM 3.3. *If the functions  $g^2$  and  $h^2$  are integrable on  $[0, 1]$  then*

$$\mathbb{E}\left(\int_0^1 g(s) dY_x(s)\right) = 0,$$

$$\mathbb{E}\left(\int_0^1 g(s) dY_x(s)\right)^2 = \sigma^2(x) \int_0^1 g^2(s) ds,$$

and

$$(23) \quad \mathbb{E}\left(\int_0^1 g(s) dY_x(s)\right)\left(\int_0^1 h(s) dY_x(s)\right) = \sigma^2(x) \int_0^1 g(s)h(s) ds.$$

*Proof.* Since  $Y_x$  is a square integrable martingale, by the Doob inequality we are able to define the above integrals in  $L^2$  sense and get the desired equalities (see e.g. Kallenberg [3]). ■

In fact the integral  $\int_0^1 g(s) dY_x(s)$  can be defined for  $g \in L^\alpha([0, 1], ds)$  in the following way:

$$(24) \quad \int_0^1 g(s) dY_x(s) = \sigma C_\alpha^{1/\alpha} \sum_{k=1}^{\infty} (\Gamma_k + x)^{-1/\alpha} \gamma_k g(U_k).$$

PROPOSITION 3.3. *If  $g \in L^\alpha([0, 1], ds)$  then the series in (24) is convergent a.s.*

*Proof.* This is an easy application of [7, Th. 4.1]. ■

REMARK 3.2. If  $g \in L^2([0, 1], ds)$  then the integral  $\int_0^1 g(s) dY_x(s)$  defined in (24) coincides (in the finite-dimensional distribution sense) with the  $L^2$  integral appearing in Theorem 3.3.

**4. Approximation of a symmetric  $\alpha$ -stable Lévy process and Brownian motion.** The process  $Y_x$  approximates the symmetric  $\alpha$ -stable Lévy process  $Z$ .

THEOREM 4.1. *The process  $Y_x$  converges uniformly on compact sets to the process  $Z$  with probability one as  $x \downarrow 0$ ; more precisely,*

$$(25) \quad \sup_{0 \leq t \leq 1} |Y_x(t) - Z(t)| \leq x\alpha^{-1} \sigma C_\alpha^{1/\alpha} Z_{\alpha/(\alpha+1)}$$

a.s. for  $x \geq 0$ , where

$$(26) \quad Z_{\alpha/(\alpha+1)} = \sum_{k=1}^{\infty} \frac{1}{\Gamma_k^{(\alpha+1)/\alpha}}$$

is an  $\alpha/(\alpha+1)$ -stable random variable with skewness parameter one and shift parameter zero.

*Proof.* Consider the function  $g(x) = a^{-1/\alpha} - (a+x)^{-1/\alpha}$  for fixed  $a > 0$  and  $x \geq 0$ . Since

$$\frac{d}{dx} [a^{-1/\alpha} - (a+x)^{-1/\alpha}] = \frac{1}{\alpha(a+x)^{1+1/\alpha}} \leq \frac{1}{\alpha a^{1+1/\alpha}}$$

it is easy to notice that

$$(27) \quad |a^{-1/\alpha} - (a+x)^{-1/\alpha}| \leq \frac{1}{\alpha a^{1+1/\alpha}} x.$$

Thus we obtain

$$\begin{aligned} & \sup_{0 \leq t \leq 1} \sigma C_\alpha^{1/\alpha} \left| \sum_{k=1}^{\infty} [\Gamma_k^{-1/\alpha} - (\Gamma_k + x)^{-1/\alpha}] \gamma_k \mathbf{I}\{U_k \leq t\} \right| \\ & \leq \sigma C_\alpha^{1/\alpha} \sum_{k=1}^{\infty} |\Gamma_k^{-1/\alpha} - (\Gamma_k + x)^{-1/\alpha}| \leq \frac{x}{\alpha} \sigma C_\alpha^{1/\alpha} \sum_{k=1}^{\infty} \frac{1}{\Gamma_k^{1+1/\alpha}} \\ & = \frac{x}{\alpha} \sigma C_\alpha^{1/\alpha} \sum_{k=1}^{\infty} \frac{1}{\Gamma_k^{(\alpha+1)/\alpha}} = \frac{x}{\alpha} \sigma C_\alpha^{1/\alpha} Z_{\alpha/(\alpha+1)} \end{aligned}$$

where  $Z_{\alpha/(\alpha+1)}$  is an  $\alpha/(\alpha+1)$ -stable random variable with skewness parameter one and shift parameter zero (the series is a.s. convergent, see Samorodnitsky and Taqqu [8]). ■

Similarly we can prove the following proposition.

PROPOSITION 4.1. *Let*

$$X(t) = \int_0^1 g_t(s) dY_x(s) = \sigma C_\alpha^{1/\alpha} \sum_{k=1}^{\infty} (\Gamma_k + x)^{-1/\alpha} \gamma_k g_t(U_k)$$

where  $g_t(s)$  is uniformly bounded. Then

$$\int_0^1 g_t(s) dY_x(s) \rightarrow \int_0^1 g_t(s) dZ(s)$$

uniformly on compact sets a.s. as  $x \downarrow 0$ .

Let us consider a symmetric  $\alpha$ -stable random vector and recall the following representation theorem (see Samorodnitsky and Taqqu [8]).

THEOREM 4.2. *If  $X$  is a symmetric  $\alpha$ -stable random vector then there are bounded measurable functions  $g_1, \dots, g_d$  such that*

$$(28) \quad X \stackrel{d}{=} \left( \int_0^1 g_1(s) dZ(s), \dots, \int_0^1 g_d(s) dZ(s) \right).$$

Thus we are able to approximate symmetric  $\alpha$ -stable random vectors by vectors with finite second moments.

THEOREM 4.3. *Let  $g_1, \dots, g_d$  be bounded measurable functions. Then*

$$(29) \quad \begin{aligned} X_x &= \left( \int_0^1 g_1(s) dY_x(s), \dots, \int_0^1 g_d(s) dY_x(s) \right) \\ &\rightarrow \left( \int_0^1 g_1(s) dZ(s), \dots, \int_0^1 g_d(s) dZ(s) \right) = X \end{aligned}$$

a.s. as  $x \downarrow 0$  and

$$(30) \quad \|X_x - X\|_d \leq \frac{x}{\alpha} M \sigma C_\alpha^{1/\alpha} Z_{\alpha/(\alpha+1)}$$

for  $x \geq 0$ , where  $\|\cdot\|_d$  is Euclidean norm,  $Z_{\alpha/(\alpha+1)}$  is the random variable defined in (26) and  $M = \sqrt{\sum_{k=1}^d \sup_{0 \leq s \leq 1} |g_k(s)|^2}$ .

*Proof.* Since the functions  $g_l$  are bounded they belong to  $L^\alpha([0, 1], ds)$  and the integral can be defined by (24). Similarly to the proof of Theorem 4.1 we can write

$$(31) \quad \left| \sum_{k=1}^{\infty} [\Gamma_k^{-1/\alpha} - (\Gamma_k + x)^{-1/\alpha}] \gamma_k g_l(U_k) \right| \leq M_l \sum_{k=1}^{\infty} |\Gamma_k^{-1/\alpha} - (\Gamma_k + x)^{-1/\alpha}| \leq M_l \frac{x}{\alpha} \sum_{k=1}^{\infty} \frac{1}{\Gamma_k^{(\alpha+1)/\alpha}} = M_l \frac{x}{\alpha} Z_{\alpha/(\alpha+1)}$$

for  $x \geq 0$ , where  $\sup_{0 \leq s \leq 1} |g_l(s)| = M_l$  and  $Z_{\alpha/(\alpha+1)}$  is defined in (26). ■



REMARK 4.1. Since  $g_1, \dots, g_d$  are bounded, they belong to  $L^2$  and by Theorem 3.3 the components of  $X_x$  have finite second moments.

The next question is what happens with the processes  $Y_x$  when  $x \rightarrow \infty$ . The answer is the following.

THEOREM 4.4.

$$(32) \quad \sqrt{\frac{2-\alpha}{\alpha\sigma^2 C_\alpha^{2/\alpha}}} x^{1/\alpha-1/2} Y_x \Rightarrow B$$

as  $x \rightarrow \infty$  weakly in the Skorokhod space  $D$  equipped with the uniform metric where  $B$  is a standard Brownian motion.

*Proof.* Since  $Y_x$  is a Lévy process, it is enough to show that  $Y_x(1) \Rightarrow B(1)$  in distribution as  $x \rightarrow \infty$  (this follows easily from [5, Th. V.19]). Taking the representation (17) of the characteristic function of  $Y_x(t)$  we find that for

$$u = u' \sqrt{\frac{2-\alpha}{\alpha\sigma^2 C_\alpha^{2/\alpha}}} x^{1/\alpha-1/2}$$

the first term in the series converges to  $tu^2/2$  and the other terms tend to zero, which yields

$$\mathbb{E} \exp\left(iu \sqrt{\frac{2-\alpha}{\alpha\sigma^2 C_\alpha^{2/\alpha}}} x^{1/\alpha-1/2} Y_x(t)\right) \rightarrow \exp\left(-\frac{tu^2}{2}\right)$$

as  $x \rightarrow \infty$ . ■

REMARK 4.2. The assertion of the above theorem can be obtained from Theorem 2.1 of [1] by substituting  $\varepsilon = \sigma C_\alpha^{1/\alpha} x^{-1/\alpha}$ .

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